The Role of Strategic and Value Chain Flexibility in Achieving Sustainability Performance: An Empirical Analysis Using Conventional and Consistent PLS

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Abstract

We contribute to the clarification of the link between dynamic and operational capabilities by examining how strategic flexibility and value chain flexibility translate into superior sustainability performance. Using survey data of chemical firms in Germany, our structural equation model shows that value chain flexibility fully mediates the relationship between strategic flexibility and sustainability performance. Further, we contribute to the ongoing research on the partial least squares (PLS) approach to structural equation modeling by estimating the proposed research model with both conventional and consistent PLS (PLSc) and outlining a guideline for evaluating and reporting PLSc-related findings.

Keywords: Strategic Flexibility, Value Chain Flexibility, Sustainability Performance, Consistent PLS, Structural Equation Modeling

1. Introduction

How companies can reconcile economic, social, and environmental performance has been actively debated in the management literature (Hart, 1995; Kolk and Pinske, 2005) While firms aim to create a differentiation-based competitive advantage by integrating sustainability into their strategy, there is a disconnection between these thoughts and action: As recently shown, only 40% of executives report that their organizations address sustainability issues and even more surprising only 10% state that their business fully tackles these issues (Kiron et al., 2013). Scholars conclude that companies should do more in order to actually transfer their direction towards sustainability issues into true business solutions (Berns et al., 2009).

One particular challenge in becoming more sustainable is to create flexibility in exploiting and controlling resources in the pursuit of alternative strategic actions. Companies that are able to flexibly allocate their resources are in a better position to deal with environmental changes, design more sustainable offerings (e.g., products that are based on renewable inputs), invest in manufacturing technologies that demand less energy or avoid toxic by-products, commit resources to new business opportunities, or reverse unproductive resource deployment (Bock et al., 2012; Zhou and Wu, 2010). Indeed, flexibility in resource allocation may explain why some companies move faster into new niches than their competitors through initiating strategic and organizational change, e.g., developing new business models, and committing resources to implement strategic actions rapidly (Eisenhardt and Martin, 2000; Nadkarni and Narayanan, 2007).

However, although strategic flexibility is considered to be indispensable to respond to environmental changes, including the increasing demand to satisfy economic needs while simultaneously considering human welfare and ecological constraints, it remains unclear how

this dynamic capability translates into superior sustainability performance (e.g., Nidumolu et al., 2009). Changing a firm's strategy might be ineffective without adequately adapting structures and processes at the operational level (e.g., Milgrom and Roberts, 1990). Helfat and Peteraf argue that "[d]ynamic capabilities do not directly affect output for the firm in which they reside, but indirectly contribute to the output of the firm through an impact on operational capabilities" (2003, p. 999). Hence, we argue that strategic flexibility - representing an important dynamic capability - does not directly impact sustainability performance (e.g., Eisenhardt et al., 2010; Teece et al., 1997). Since strategic flexibility rather builds, integrates, and reconfigures operational capabilities, we eventually propose value chain flexibility as an important operational capability in the pursuit of superior sustainability performance (Helfat and Peteraf, 2003; Kortmann et al., 2014; Teece et al., 1997). Value chain flexibility relates to the firm's ability to adapt its operational activities across the entire value chain, including product development, manufacturing, logistics, and spanning activities, to changing customer needs (Zhang et al., 2002).

Taken together, this study empirically examines the mediating role of value chain flexibility within the relationship of strategic flexibility and sustainability performance. Using a top-level manager dataset of chemical firms located in Germany, we contribute to the ongoing debate on the triple bottom line within the management literature by showing how superior sustainability performance results from a firm's ability to (i) flexibly allocate its resources at the strategic level and (ii) adequately adapt it structures and processes at the operational level. Furthermore, we contribute to the ongoing research on the partial least squares (PLS) approach to structural equation modeling (SEM). Following most recent literature, we estimated our proposed research model by means of consistent PLS (PLSc) as an extension to conventional PLS. While maintaining all strengths of conventional PLS, PLSc overcomes the lack of consistency and corrects for attenuation of regression path estimates. We finally examine whether the results deriving from conventional PLS and PLSc significantly differ for our proposed research model. In so doing, we borrow from the literature on PLS-SEM and outline a guideline for evaluating and reporting PLSc-related findings.

2. Hypotheses

2.1 Strategic Flexibility and Value Chain Flexibility

As business environments have become more competitive and dynamic than ever before, companies are continuously forced to adapt to environmental changes (Grewal and Tansuhaj, 2001; Young-Ybarra and Wiersema, 1999). The ability of a company to rapidly identify major changes in the competitive landscape, reallocate resources to new courses of action and reconfigure existing organizational routines that support these actions, will ultimately determine whether a company can faster create competitive advantage than its rivals (e.g., Eisenhardt, 1989; Shimizu and Hitt, 2004; Nadkarni and Narayanan, 2007). This adaptive capability is known as strategic flexibility and refers to the "ability of a firm to reallocate and reconfigure its organizational resources, processes, and strategies to deal with environmental changes" (Zhou and Wu, 2010, p. 549). Following prior literature, strategic flexibility is classified as an important dynamic capability (Eisenhardt et al., 2010; Teece et al., 1997), which represents the ability to build, integrate, and reconfigure operational capabilities (Helfat and Peteraf, 2003; Kortmann et al., 2014; Teece et al., 1997). Hence, by continuously creating and recombining resources in novel ways, dynamic capabilities have a direct impact on operational capabilities (Helfat and Peteraf, 2003; Kortmann et al., 2014). The latter involve the execution and coordination of various tasks, such as the development, manufacturing, and marketing of products and services (Helfat and Peteraf, 2003). Based on this classification of dynamic and operational capabilities, we argue that strategic flexibility is positively associated with value chain flexibility. Value chain flexibility represents an operational capability that ensures the flexible adaption of operational activities across the entire value chain, including product development, manufacturing, logistics, and spanning activities, to changing customer needs (Zhang et al., 2002). Strategic flexibility enables companies to redeploy assets such as product-creating resources to meet a variety of customer expectations (e.g., meeting high ecological standards while reducing product price) and supplier demands without performance loss. Through enabling response to unanticipated changes and unexpected consequences of predictable changes (Bahrami, 1992), strategic flexibility acts as an organizing principle, which enhances the coordination of product design, production and distribution, and allows firms to take operational actions, e.g., adjust production volume or respond to ad-hoc requests made by customers (Nair, 2005; Young-Ybarra and Wiersema, 1999; Zhang et al., 2003). We therefore hypothesize:

H1: Strategic flexibility is positively associated with value chain flexibility.

2.2 Value Chain Flexibility and Sustainability Performance

In order to effectively and efficiently respond to changing customer needs, including the increasing demand for sustainable offerings, firms are forced to re-design their operational processes and implement a flexible system of structures and processes across their value chain (e.g., Nair, 2005; Kolk and Pinske, 2005; Zhang et al., 2002; Zhang et al., 2003). Nidumolu et al., for instance, argue that operational innovations are central to building a sustainable value chain and, hence, conclude that companies need to develop the "ability to re-design operations" (2009, p. 61). Following Zhang et al., "the breadth and intensity of flexibility needed to cope with changing customer requirements [however] cannot be provided by one department or function" (2003, p. 175). In order to increase responsiveness and to eliminate bottlenecks across the value chain, flexibility in operational activities needs to be present in various activities, including product development, manufacturing, logistics, and spanning activities (Zhang et al., 2002; Zhang et al., 2003). These activities include the integration, coordination, and communication across the internal as well as external value chain (e.g., Zhang et al., 2003). Considering that the increasing demand for sustainable solutions very often implies a shift from the sole focus on end-customers toward the consideration of earlier stages of the value chain, the ability to share information with suppliers and customers is indispensable in the pursuit of superior sustainability performance. Although sustainability has been identified as an emerging global trend in almost all industries, the willingness to pay an extra price for sustainable offering still varies among customers. Hence, the successful introduction of sustainable products or services might also strongly benefit from the firm's ability to externally and internally share both standardized (e.g., information on the general function and usability of a product or service) and customized information (e.g., information on customer-specific needs associated with sustainability). Based on this discussion, we propose:

H2: Value chain flexibility is positively associated with sustainability performance.

3. Methodology

3.1 Sample and Key Informant Check

The data of the present study derive from an undertaken online survey of top-level managers of chemical firms located in Germany. The chemical industry was chosen as research context since it reflects an industry that has been successful in making business cases for sustainability (e.g., Jenck et al., 2004). For data collection, we choose an online questionnaire with closed questions. Apart from advertising our study within a practitioner-oriented magazine, we directly invited top-level managers via email to participate in our online survey.

3.2 Non-Response Bias

Since the present study might be potentially threatened by non-response bias, we performed a post-hoc analysis and tested for differences between (i) early and late respondents (Armstrong and Overton, 1977; Li and Calantone, 1998) and (ii) participants that completed the survey and participants that abandoned. Both Mann-Whitney U-tests revealed no significant differences between the different groups.

3.3 Key Information Check

For data collection, we applied a key informant approach and solely included responses of top-level managers in the final dataset of our study. To reduce the potential risk that participants respond to questions beyond their level of responsibility, we applied the following key informant criteria: (i) involvement in strategic, innovation, and operational decision making, (ii) job title, (iii) job experience, and (iv) organizational tenure (see Appendix A).

3.4 Measures

Our proposed research model consists of three multi-item reflective scales. Strategic flexibility was measured using a six-item construct adapted from Zhou and Wu (2010). Value chain flexibility was measured using a six-item construct adapted from Nair (2005), who drew upon the conceptualization introduced by Zhang et al. (2002). Sustainability performance is a new performance measure we developed to operationalize competitive advantage that is achieved by meeting the needs of organizational stakeholders while simultaneously considering human welfare and ecological constraints (e.g., Chow and Chen, 2012; Nidumolu et al., 2009). The measurement items are outlined in Appendix B.

3.4 Common Method Variance

In order to test whether our findings are potentially threatened by common method bias, we included a common method variance factor that comprises all principal constructs' indicators in our structural model (Podsakoff et al., 2003, Liang et al., 2007). We compared the variance of each principal construct's indicator as explained by the common method and substantive factors. The model estimation with PLS revealed an average substantive explained variance of 0.71 and an average common method based variance of 0.01. The resulting ratio of the average substantive explained variance to the average common method based variance is about 71:1.

4. Analysis and Results

Following recent research across various business disciplines, e.g. marketing, operations management, and information systems, we applied PLS-SEM. Among variance-based SEM techniques, PLS can be regarded as the most advanced approach to SEM (Dijkstra and Henseler, 2015a). PLS path modeling is to be preferred over alternative covariance-based techniques, when the primary aim of the study rather relates to causal-predictive analysis than theory testing (e.g., Hair et al., 2012; Henseler et al., Sinkovics, 2009). Compared to covariance-based techniques, PLS-SEM has minimum demands regarding sample size (Chin et al., 2003; Hair et al., 2012). Furthermore, PLS-SEM is more favorable when the proposed research model comprises a high degree of complexity (Chin and Newsted, 1999; Henseler and Chin, 2010). However, similar to other analysis techniques, PLS-SEM is not without disadvantages. The main downsides of PLS-SEM are the lack of (i) consistency and (ii) and overall goodness-of-fit measure (e.g., Dijkstra and Henseler, 2015b; Henseler and Sarstedt, 2013). PLS estimates are only consistent under the conditions of 'consistency at large' and, hence, demand large numbers of both observations and measurements items per construct (Dijkstra and Henseler, 2015a,b; Peng and Lai, 2012). The reason is the approximation of latent variables as linear

combinations of their respective observed indicators within PLS-SEM (Henseler et al., 2014; Dijkstra and Henseler, 2015b). While this approach is proper to estimate composite factor models, the PLS-SEM is not consistent for common factor models (e.g., Henseler et al., 2014; Dijkstra and Henseler, 2015a,b). Considering that the PLS algorithm approximates common factors as linear composites of observed indicators (Dijkstra and Henseler, 2015a,b), prior scholars have acknowledged that "parameter estimates for paths between observed variables and latent variable proxies are biased upward in PLS (away from zero), while parameter estimates for paths between proxies are attenuated" (Gefen et al. 2011, p. vi). To overcome the lack of inconsistency, recent literature has introduced an important advancement to PLS: consistent PLS (PLSc) (Dijkstra and Henseler, 2015a,b). PLSc corrects the estimates of reflectively measured constructs deriving from the traditional iterative PLS algorithm by employing a new reliability coefficient, termed as Dijkstra-Henseler's rho (pA). Considering the advantages implied by PLSc, we employed the PLSc approach to SEM in addition to conventional PLS, using SmartPLS 3.0 (Ringle, Wende, and Becker, 2015) and Adanco 1.0 (Composite Modeling, 2014). Borrowing from the literature on PLS-SEM (e.g., Gefen et al., 2011; Hair et al., 2012; Henseler et al., 2009; Hulland, 1999; Peng and Lai, 2012), we outline a guideline for evaluating and reporting PLSc-related findings.

4.1 Validity and Reliability

Following prior literature (e.g., Hair et al., 2011; Hulland, 1999), we assessed the outer model by means of (i) individual item reliability, (ii) internal consistency reliability, (iii) convergent validity, and (iv) discriminant validity. Individual item reliability refers to the correlation of each indicator with its latent variable and, hence, is represented by the outer loadings within PLS/PLSc. While all items exceed the suggested threshold of 0.70 with regard to the estimation using PLS (Henseler et al., 2009; Hulland, 1999), the estimation of our proposed research model with PLSc reveals some outer loadings that are below the recommended cut-off value of 0.70 (see Table 1). The comparison between PLS and PLSc in Table 1 eventually supports prior literature (Dijkstra and Henseler, 2015a), which also shows that the estimation with conventional PLS leads to an overestimation of the absolute value of loadings.

Table 1. Individual item reliability (outer loadings)

		tegic bility		Chain bility	Sustainability Performance			
	PLS	PLSc	PLS	PLSc	PLS	PLSc		
SF_1	0.76	0.58						
SF_2	0.80	0.76						
SF_3	0.74	0.49						
SF_4	0.91	0.96						
SF_5	0.84	0.90						
VCF_1			0.75	0.68				
VCF_2			0.78	0.67				
VCF_3			0.84	0.90				
VCF_4			0.90	0.77				
VCF_5			0.84	0.84				
VCF_6			0.72	0.71				
SP_1					0.90	0.90		
SP_2					0.92	0.88		
SP_3					0.90	0.94		
SP_4					0.88	0.80		
<u>SP_5</u>					0.88	0.81		

Furthermore, we evaluated the internal consistency reliability of our measurement constructs. While Cronbach's alpha (CA) and composite reliability (CR) are commonly applied criteria for the evaluation of reflective measurement constructs, prior scholars suggest to prefer

CR over CA when using PLS-SEM (Chin, 2010; Hair et al., 2012). The limitation of CA is grounded in its bias against short measurements scales and the assumption that all indicators are equally reliable (; Hair et al., 2011). Thus, we specifically emphasize to consider these limitations when interpreting CA estimates using PLS-SEM. For adequate internal consistency reliability, values of 0.70 or higher are suggested for both CA and CR (Nunnally, 1978). As shown in Table 2, all measurement constructs exceed the suggested threshold of 0.70 and, hence, indicate sufficient internal consistency reliability. However, as Dijkstra and Henseler (2015a, p. 3) state, "Cronbach's coefficient alpha (Cronbach 1951) and composite reliability (Chin 2010), are not consistent themselves". Since CA tend to underestimate and CR tend to overestimate the actual reliability of construct scores, the authors recommend to refrain from using CA and CR when estimating models with PLSc (Dijkstra and Henseler, 2015a). Instead, researchers should evaluate and report construct reliability by means of the Dijkstra-Henseler's rho (p_A) reliability coefficient (Dijkstra and Henseler, 2015a). Similar to CA and CR, the reliability coefficient p_A should indicate values of 0.70 or higher in exploratory research and values above 0.80 or 0.90 for more advanced stages of research (Hair et al., 2011; Henseler et al., 2009; Nunnally and Bernstein, 1994). As shown in Table 2, the reliability coefficient ρ_A of each measurement construct is above 0.70.

Convergent validity was evaluated by means of the average variance extracted (AVE). Prior literature suggests an AVE value of 0.50 or higher (Hair et al., 2011; Henseler et al., 2009). The estimation with both conventional PLS and PLSc reveal satisfactory AVE values with regard to all measurement constructs (see Table 2). The comparison of both estimation techniques eventually shows that the AVEs deriving from PLSc are slightly below those deriving from traditional PLS. Moreover, we assessed the outer model by means of discriminant validity. Here, we firstly referred to the Fornell-Larcker criterion, which "postulates that latent variable shares more variance with its assigned indicators with another latent variable in the structural model" (Hair et al., 2011, p. 146). In order to fulfill this criterion, the square root of each AVE should be greater than its correlation coefficients with any other latent construct. The Fornell-Larcker criterion is fulfilled with regard to all measures (see Table 2). Secondly, we assessed discriminant validity on the indicator level. To this end, we referred to the crossloadings and evaluated whether each indicator loading with its associated construct exceeds is loading with each of the other constructs (Hair, Ringle, and Sarstedt, 2011; Henseler, Ringle, and Sinkovics, 2009). As shown in Appendix B, this criterion is fulfilled with regard to all measurement items.² Thirdly, we drew upon the recent literature (Henseler, Ringle, and Sarstedt, 2015) and assessed discriminant validity by means of the heterotrait-monotrait ratio (HTMT). While some scholars propose a threshold of 0.90 (e.g., Teo et al., 2008), others recommend values of HTMT that are below 0.85 in order to ensure discriminant validity (Clark and Watson, 1995; Kline, 2011). The HTMT values and the corresponding confidence intervals deriving from a bootstrapping procedure with 500 resamples are outlined in Table 3.3 Since none of our measurement constructs violates the HTMT_{0.90} and HTMT_{0.85} criterion and all constructs additionally fulfill the HTMT_{inference} test, we infer adequate discriminant validity (Henseler, Ringle, and Sarstedt, 2015).

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¹ For a more detailed description of the new consistent reliability coefficient ρ_A, please refer to Dijkstra and Henseler, 2015a).

² The indicator loadings indicated in Appendix B derive from the conventional iterative PLS algorithm. Since the existing software packages (e.g., SmartPLS 3.0 or Adanco 1.0) do correct the indicators' cross-loadings with other constructs, discriminant validity on the indicator level was not evaluated with regard to the model estimation using PLSc.

³ The calculation of HTMT is based on the correlations of individual indicators (Henseler et al., 2015). Since the estimation with PLS and PLSc leads to the same item correlation matrix, the values of HTMT do not differ with regard to PLS and PLSc. Hence, the HTMT values and their corresponding confidence intervals presented in Table 3 relate to the model estimation with both PLS and PLSc

Table 2. Properties of measurement scales and correlations

	ME	SD	CA	CR (p _c)	$\mathbf{p}_{\mathbf{A}}$	AVE	1	2	3	4	5
Strategic Flexibility	4.45	1.20	0.87 (/)	0.91 (/)	/ (0.91)	0.66 (0.58)	0.81 (0.76) ^a				
2. Value Chain Flexibility	4.54	1.33	0.89 (/)	0.92 (/)	/ (0.90)	0.65 (0.58)	0.47 (0.52)	0.81 (0.76) ^a			
3. Sustainability Performance	4.38	1.49	0.94 (/)	0.95 (/)	/ (0.94)	0.80 (0.75)	0.21 (0.23)	0.47 (0.51)	0.89 (0.87) ^a		
4. Firm Age	86.10	53.68	- (-)	- (-)	- (-)	- (-)	0.10 (0.11)	0.09 (0.09)	0.36 (0.37)	-	
5. Firm Size	4.53	1.24	- (-)	- (-)	- (-)	- (-)	-0.03 (-0.03)	0.01 (0.01)	0.17 (0.17)	0.63 (0.63)	-

Note: ME = Mean. SD = Standard Deviation. CR = Composite Reliability. CA = Cronbach's Alpha. p_A = Dijkstra-Henseler's rho. AVE = Average Variance Extracted. ^a Value on the diagonal is the square root of AVE. Values indicated as follows: PLS (PLSc).

Table 3. Heterotrait-monotrait ratio (HTMT)

	Strategic Flexibility			Value C	Chain Fl	exibility		stainability erformance		Firm Age		
	нтмт		dence rval	нтмт		dence rval	нтмт	Confidence Interval		нтмт	Confidence Interval	
		Low	Up		Low	Up		Low	Up		Low	Up
Value Chain Flexibility	0.51	0.31	0.69									
Sustainability Performance	0.23	0.00	0.44	0.50	0.28	0.69						
Firm Age	0.11	-0.12	0.32	0.09	-0.14	0.30	0.37	0.15	0.55			
Firm Size	-0.04	-0.25	0.21	0.01	-0.19	0.21	0.17	-0.11	0.42	0.63	0.51	0.73

4.2 Results

While the path coefficients were estimated with the path method (Chin, 2001), their corresponding standard errors have been estimated by performing a bootstrapping procedure with replacement, using 500 resamples (Chin, 1998; Nevitt and Hancock, 1998). Figure 1 and Table 4 entail the results we obtained using both PLS and PLSc.

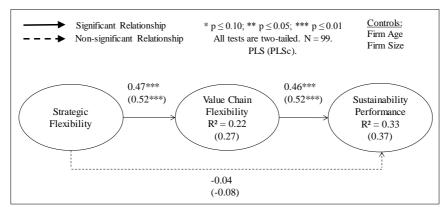


Figure 1. Results of structural equation modeling with PLS and PLSc

Table 4. Results of structural equation modeling with PLS and PLSc

	Value Ch	ain Flexibility	Sustainabilit	y Performance
	β-value	p-value	β-value	p-value
Controls				
Firm Age	0.04 (0.03)	0.75 (0.83)	0.37 (0.38)	< 0.001 (< 0.001)
Firm Size	0.00 (0.01)	0.99 (0.96)	-0.07 (-0.07)	0.62 (0.61)
Main Effects				
Strategic Flexibility	0.47 (0.52)	< 0.001 (< 0.001)	-0.04 (-0.08)	0.71 (0.53)
Value Chain Flexibility			0.46 (0.52)	<0.001 (<0.001)
R-Square	0.22 (0.27)		0.33	3 (0.37)
All tests are two-tailed.	N = 99. PLS (1	PLSc)		

4.3 Mediation Analysis

To further define the degree of the mediating effect of value chain flexibility within the strategic flexibility-sustainability performance relationship, we followed Subramani (2004) and compared two competing models: (i) a full mediation model (research model) and (ii) a partially

mediated model (nested model). We specifically assessed whether the incorporation of a direct path substantively improves the explanation of the dependent variable. As shown in Table 5, value chain flexibility fully mediates the relationships between strategic flexibility and sustainability performance with regard to the estimation with both PLS ($f^2 = 0.001$; p = 0.708) and PLSc ($f^2 = 0.003$; p = 0.584).⁴

Table 5. Nested-model comparison

Estimation Method	R ² in Nested Model	R ² in Full model	Magnitude of the change (f^2)	Pseudo F	p-value	Conclusion ^a			
PLS	0.325	0.326	0.001	0.141	0.708	Full mediation			
PLSc	0.368	0.370	0.003	0.302	0.584	Full mediation			
^a Significance level: p ≤ 0.05.									

4.4 Model Fit and Prediction Analysis⁵

To assess the prediction power of our proposed research model, we referred to the explained variance (R²) of all dependent variables. Since our results indicate a R² value of 0.22 (0.27) for value chain flexibility and 0.33 (0.37) for sustainability performance, our proposed research models shows weak to moderate prediction power (Chin, 1998). We also evaluated the effect size of the predictor constructs by computing their corresponding Cohen's f² values (Cohen, 1988; Henseler et al., 2009). Cohen's f² values are represented in Table 6. As rule of thumb, prior scholars (Cohen, 1988; Chin, 1998) recommend Cohen's f² values of 0.02, 0.15, and 0.35 for small, medium, and large effect size. Furthermore, we evaluate the prediction relevance of our proposed research model by means of a non-parametric Stone-Geisser test. The Stone-Geisser Q^2 values derive from a blindfolding procedure with an omission distance of 5 and were calculated by means of the cross-validated redundancy approach (Hair et al., 2012). Since all values are greater than zero, we infer that all endogenous constructs show adequate predictive relevance (see Table 6) (Götz et al., 2010; Henseler et al., 2009). Table 6 further indicates the relative prediction relevance (q^2) of each predictor variable for the endogenous constructs of our proposed research model. The effect of strategic flexibility on value chain flexibility ($q^2 = 0.159$) and the effect of value chain flexibility on sustainability (q^2 = 0.179) both show small to medium prediction relevance.⁶

Table 6. Effect size (Cohen's f^2) and relative prediction relevance (q^2)

		Effect size		Relative prediction relevance				
Path	R_{incl}^2	R_{excl}^{2}	Cohen's f ²	Q_{incl}^2	Q_{excl}^2	q^2		
Strategic Flexibility – Value Chain Flexibility	0.223 (0.273)	0.011 (0.012)	0.273 (0.359)	0.140 (/)	0.003 (/)	0.159 (/)		
Strategic Flexibility – Sustainability Performance	0.326 (0.373)	0.325 (0.368)	0.001 (0.008)	0.255 (/)	0.255 (/)	0.000 (/)		
Value Chain Flexibility – Sustainability Performance	0.326 (0.373)	0.166 (0.178)	0.237 (0.311)	0.255 (/)	0.122 (/)	0.179 (/)		
Values presented as follows: PLS (PLSc).								

In contrast to CB techniques for SEM, PLS lacks a global goodness-of-fit measure. To overcome this shortcoming, prior scholars (e.g., Vinzi et al., 2010; Tenenhaus et al., 2005) suggest the use of the goodness-of-fit' (GoF) criterion, which can be understood as the geometric mean of the average communality and average explained variance. Both model estimation with PLS reveal a GoF that exceeds the suggested threshold of 0.36 for large effect

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⁴ Moreover, we evaluated the magnitude and significance of the indirect path between strategic flexibility and sustainability performance via value chain flexibility. For both estimation techniques, i.e., PLS (β = 0.21, p < 0.001) and PLSc (β = 0.21, p < 0.001), the indirect path is positive and significant.

⁵ Values presented as follows: PLS (PLSc).

⁶ Since PLSc does not impact the consistency of the predictive relevance of the research model, we solely performed the blindfolding procedure for the model estimation with PLS.

sizes of R² (PLS: 0.44; PLSc: 0.45). However, more recent literature particularly shows that the measure is not suitable for model validation (Henseler and Sarstedt, 2013). Hair et al. (2014), therefore, recommend evaluating the model's overall quality in terms of how well it predicts the endogenous constructs. Considering the inadequacy of GoF as global goodness-of-fit measure, we followed Henseler et al. (2014) and additionally referred to the standardized root mean square residual (SRMR) as index for model validation. The absolute measure of model fit is defined as the difference between the observed correlation and the predicted correlation. Values below 0.08 are generally considered as favorable (Hu and Bentler, 1999). While the model estimation with PLS (composite factor model) reveals a SRMR value of 0.075, the estimation with PLSc (common factor model) indicates a SRMR value of 0.164.

5. Discussion and Implications

Product innovations that transform the nature of production and consumption, an increasingly interconnected global economy that is altering social conditions, and an environment that is more unpredictable than ever before, have prompted sustainability as an important business issue on the agenda of many managers and force them to question existing assumptions of how to create value (Kiron et al., 2013). As Nidumolu et al. argue, "the quest for sustainability is already starting to transform the competitive landscape, which will force companies to change the way they think about products, technologies, processes, and business models" (2009: 58). Considering sustainability as an important business issue, our study underlines the critical role of flexibility in both strategic and operational actions in achieving high sustainability performance. Stimulated by strategic flexibility, value chain flexibility enables firms to deliver sustainable products and services in a timely manner, cope with changing customer requirements, shorten lead times and reduce inventory costs (Zhang et al., 2002). Apart from contributing to the ongoing debate on the triple bottom line within the management literature, these findings contribute to the dynamic-resource based view of the firm (e.g., Helfat and Peteraf, 2003; Teece, 1997). The results of our empirical analysis demonstrate that an operational capability, here value chain flexibility, fully mediates the relationship between a dynamic capability, here strategic flexibility, and sustainability-related performance outcomes. Hence, our empirical findings reveal strategic flexibility and value chain flexibility as distinct yet interlinked capabilities in the pursuit of superior sustainability performance. By this means, we empirically support Helfat and Peteraf, who argue that "[d]ynamic capabilities do not directly affect output for the firm in which they reside, but indirectly contribute to the output of the firm through an impact on operational capabilities" (2003, p. 999).

Furthermore, we contribute to the ongoing research on the PLS approach to SEM by estimating our proposed research model using both conventional PLS and consistent PLS. The comparison of both estimation methods reveals the typical pattern of results with regard to traditional PLS: While the outer loadings were overestimated, the inter-construct correlations were attenuated (Dijkstra and Henseler, 2015b). Considering the high values of the reliability coefficient ρ_A , the subsequent correction undertaken by PLSc however was rather weak, causing merely small differences the estimates deriving from PLS and PLSc. Both estimation methods eventually lead to the same conclusion: value chain flexibility fully mediates the relationship between strategic flexibility and sustainability performance.

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Appendix A: Descriptive Statistics

Key Informant Descriptive Statistics		Firm Descriptive Statistics					
Job Title		Firm Size (Number of l	Full Time Employees)				
CEO	15	1-10	3				
CTO	1	11-50	8				
Vice President	1	51-250	6				
Executive Director	1	251-1000	14				
Director	4	1.001-50.000	53				
Chairman	2	> 50.000	15				
Business Unit Manager	5	_	AVG (SD)				
Head of Department	21	Firm Age (in years)	86.10 (53.68)				
Senior Manager	2						
Partner	1						
Operations Manager	6						
General Manager	40						
Involvement in	AVG (SD)	,					
strategic decision making	5.14 (1.72)						
innovation decision making	5.39 (1.52)						
operation decision making	4.91 (1.77)						
Organizational tenure (in years)	12.64 (9.87)						
Overall work experience (in years)	19.87 (10.23)						
Age (in years)	46.23 (11.35)						

Appendix B: Measures

	ME	SD	1	2	3
1. Strategic Flexibility (Zhou and Wu, 2010)					
SF 1 The flexible allocation of marketing resources (including advertising, promotion and distribution resources) to market a diverse line of products.	4.36	1.37	0.76	0.24	0.24
SF 2 The flexible allocation of production resources to manufacture a broad range of product variations.	4.39	1.58	0.80	0.38	0.16
SF 3 The flexibility of product design (such as modular product design) to support a broad range of potential product applications.	4.39	1.53	0.74	0.25	0.10
SF 4 The redefinition of product strategies in terms of target market segments.	4.70	1.47	0.91	0.48	0.21
SF 5 The reallocation of organizational resources to support the firm's intended product strategies.	4.34	1.42	0.84	0.47	0.15
2. Value Chain Flexibility (Nair, 2005)					
VCF 1 We have a flexible program of special services that can be matched to changing customer requirements.	4.29	1.70	0.41	0.75	0.28
VCF 2 We have established a program to authorize and perform special requests made by selected customers	4.89	1.65	0.40	0.78	0.26
VCF 3 We are able to accommodate a wide range of unique customer requests by implementing pre-planned solutions.	4.36	1.70	0.42	0.84	0.47
VCF 4 We have adequate ability to share both standardized and customized information externally with suppliers and/or customers.	4.54	1.66	0.36	0.90	0.39
VCF 5 We have adequate ability to share both standardized and customized information internally.	4.87	1.57	0.34	0.84	0.49
VCF 6 We have increased operational flexibility through supply chain collaboration.	4.20	1.61	0.36	0.72	0.34
3. Sustainability Performance					
SP 1 We are the first that offer environmental-friendly products/services at the marketplace.	4.37	1.62	0.21	0.43	0.90
SP 2 Our competitors consider us as a leading company in the field of sustainability.	4.09	1.67	0.18	0.44	0.92
SP 3 We develop new products/services or improve existing products/services that are regarded as sustainable for society and environment.	4.59	1.79	0.20	0.41	0.90
SP 4 Our reputation in terms of sustainability is better than the sustainability reputation of our competitors.	4.48	1.68	0.16	0.40	0.88
SP 5 Compared to our competitors, we more thoroughly respond to societal and ethical demands.	4.37	1.55	0.21	0.42	0.88
Notes: ME = Mean; SD = Standard Deviation.					