An empirical assessment of the impact of public research contracts on scientific productivity*

César Alonso-Borrego[†] Auniversidad Carlos III

Antonio Romero-Medina Universidad Carlos III

Rocío Sánchez-Mangas Universidad Autónoma de Madrid

Abstract

We study the effect of a Program, aimed at financing high quality researchers to integrate them into the Spanish scientific system, on the relative ex post performance of the researchers awarded. We assess the effect of the contract status on the scientific productivity of applicants, in the four-year period after application. Both the conditional regression and the matching results show that the contract status has no effect on the average number of published contributions, but it exhibits, for several areas, a positive effect on the scientific quality of contributions (as measured by its impact). This result points out the success of the Program in increasing the scientific impact of the Spanish system.

Keywords: Brain Gain, Government Research Programs, Human Capital, Policy Evaluation.

JEL Classification Numbers: O38, D78, C21, C78.

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[†]Corresponding Author. Department of Economics, Universidad Carlos III de Madrid, 28903 Getafe, Spain. e-mail: cesar.alonso@uc3m.es.

1 Introduction

This paper studies if the Ramon y Cajal (R&C) contract has a differential impact on the research output of the researchers awarded with Ramon y Cajal contracts years later. For such purpose, we exploit data on applications in several calls of the Program, as well as individual and curricular information of the applicants.

The Ramon y Cajal Program was introduced in 2001 by the Spanish Government. It is an ambitious publicly funded program aimed at providing career paths to high-quality researchers and to integrate them within the national scientific system. It was created in the general context of a lack of R&D personnel in Spain and with Spanish Universities hiring policies being called into question.

The researchers hired under the program benefitted from a well-defined career path, with a 5-year contract in a Spanish research centers access to research grants, and the possibility to join permanent research positions at the end of the contract (see Sanz Menéndez et al, 2002; Sanz Menéndez, 2003).

The selection procedure was centralized in an evaluation agency, Agencia Nacional de Evaluación y Prospectiva (ANEP). This agency, appointed by the Government, appraised all eligible applicants based on predetermined rigorous and objective evaluation criteria based mainly on the candidates' scientific record. For this purpose, 24 evaluation committees of national

and international experts, one for each research field, were constituted by the evaluation agency.¹ Four years after each call, the performance of each granted applicant during the contractual period was evaluated. A positive evaluation implies the possibility to receive a new contract that facilitated her access to a tenured contract in the research center.

We confirm that the assessment process was based on the applicant's research C.V. and that the available curricular information has mattered for the grading that the assessment committees gave to each applicant. In order to assess the effectiveness of the Program, we analyze the impact of the contract status on researchers' scientific productivity four years after the call. To circumvent potential selection biases due to differences between successful and unsuccessful applicants, we control for observed curricular characteristics that yield a similar probability of contract at the time of the call. Our main results show that the contract had no impact on the number of published contributions, but had a positive effect on the quality of contributions.

The rest of the paper is organized as follows. In section 2, we introduce the main data set of application, and the complementary data set on the applicants' curricular information and preliminary results. In section 3, we evaluate the effectiveness of the Program in the scientific productivity of successful applicants. In section 4, we summarize the major results and discuss their policy implications, and conclude.

¹A list of the 24 research areas is shown in the Appendix, Table A1.

2 Data and preliminary evidence

The main data set, provided by the Dirección General de Investigación of the Spanish Ministry of Education, records all applications in the first seven calls of the program, from 2001 to 2007. We excluded observations with missing values for individual characteristics, which represent less than one percent of all observations. Each applicant information includes her research area, the institution and year when she earn her PhD, country of residence and nationality, as well as the score received in the assessment process and whether she was granted a contract.

In Table 1, we provide the distribution of all the applications and the distribution of successful applications in Table 1. The distribution is not uniform in time, as it is not the number of contracts offered, which decreased substantially since 2003. In addition, since 2004, eligibility was restricted to earning the PhD in the last 10 years, and a minimum 2-years stay in center different than that in which the Ph.D. was obtained.

We observe that applications are dominated by men. The research areas of Biology, Chemistry and Medicine cope about 60% of the applications. It must be noted that the gender distribution is strongly unequal across research areas. Physics and Engineering are strongly dominated by men, amount 80% of applicants. In Chemistry and Business, men represent about 60%. However, in Social Sciences and in Biology, the proportion of men is around 52%. Medicine is dominated by women, with 54% of applications.

With respect to time elapsed since the PhD, the majority of applicants earned their Ph.D. within 3 to 6 years before the call.

The curricular information has been collected from a complementary data source, the free net resource Publish or Perish (Harzing, 2007). Publish or Perish retrieves academic contributions by author using the Google Scholar database, which provides the title, the source, the year and the authors of the contribution. Google Scholar is generally praised for its speed (Bosman et al., 2006) and its high correlation with alternative bibliometric sources (See Harzing, 2012, and Harzing and van der Wal, 2011, for a comparison of citation analysis using different data sources). Whenever the contribution was published in a scientific journal, the journal information is also reported. For each applicant, we can then measure her number of distinct contributions and, among these, the number of published papers. In order to weight the quality of each contribution, we use the Journal of Citation Reports (JCR), which provides the impact factors of the international journals listed in its database. The impact factor of a journal is calculated on the basis of the average number of citations attained by the contributions published in that journal. We use the JCR impact factors to measure both the quality of each candidate, as well as the quality of the center where each candidate earned her PhD, defined as the average number of citations to all the works published in JCR journals by all the researchers affiliated to the center. The curricular information is updated up to 2008.

We concentrate on the first calls, until 2003, as we have to exploit curricu-

lar information several years after the call. Data on 2001 is also disregarded, given the special characteristics of the first call, which might harden the comparability of applications with subsequent calls (Alonso-Borrego et al, 2013). Moreover, we have restricted the sample, for the sake of comparability, to applicants who earned a PhD within the last ten years.² Our final sample, therefore, contains 4,967 applicants between 2002 and 2003.

We use three measures of scientific quality of each applicant: her number of contributions listed in the JCR database, the average impact factor of her JCR publications, and the maximum impact factor among the JCR journals in which she has published. The two impact factor measures are based on the corresponding impact of the journal in which she has published each contribution.

In Table 2 we summarize the curricular information of applicants by contract status. Besides, we break down the sample by applicants' characteristics: gender, research area, and time elapsed since the PhD. For all categories considered, we observe that, at the time of the call, granted researchers have, on average, more published contributions and a higher scientific impact (either average or maximum impact) than non-granted researchers. Nevertheless, given the high standard deviations, most differences are not significant. We also find that the three measures of scientific quality differ substantially by area, reflecting differences in the usual number of papers and citations among areas.

²Actually, this condition was established by the Program since 2004.

To ascertain the factors that are relevant for the committees' assessment of applicants, we consider a conditional analysis of the applicants' scores on their individual and curricular characteristics. The OLS estimation results are shown in Table 3. The first column reports the full sample regression of score on applicant's characteristics, using qualitative variables to allow for differences among areas. We also report separate estimations for each of the Publish or Perish areas.³ In all the estimates, both for the full sample and by areas, a high proportion of the variance of score is explained. In most areas, scientific quality of applicants prove to be determinant in the committees' assessments, and in general the quality of contributions matters more than its quantity. Furthermore, the quality of the center in which the Ph.D. was obtained is also a relevant factor in Physics, Economics and Humanities. In the case of Economics and Humanities, the curricular variables are non significant. The small sample size in Economics lead to very imprecise estimated coefficients. This is not the case, though, of Humanities, what suggests that the committee's assessment relies on different criteria. In Physics, the number of JCR papers appears more relevant than the quality of contributions. It is also interesting that PhD tenure has an inverted-U effect on applicants' score in most areas. Given that we are also controlling for scientific quality, this variable, together with the quality of the PhD center, might be capturing other unobserved quality features. For instance: papers on under revision;

³A list of the 7 areas reported in Publish or Perish is shown in the Appendix, Table A2.

forthcoming papers (but not published at the year of the call); the quality of the research agenda of the candidate, etc.

Among the individual characteristics, we find a positive and significant gender effect in favor of men, which ranges between 3 and 5 percentage points. This result suggests that men are slightly better graded than women with similar scientific quality, pointing out a certain degree of gender discrimination. Also, we have included the quality of the PhD institution, measured by the cumulated impact of the contributions listed in the JCR of its faculty members, which has a positive effect, and it is significant in several areas.

3 The performance of Ramon y Cajal researchers

We are mostly intrigued about the impact of the contract status on the ex-post performance of researchers. For that purpose, we consider the scientific outputs of applicants in the four years after the call. Given the data constrains, we consider the time horizon chosen to be sufficient to test the potential influence of the contract. It is, though, consistent with the usual time span for the tenure decision taken by the research centers. Moreover, such time span seems to be in coherence with the maximum time length needed to undertake a peer-reviewed publication process of scientific contributions.

Our relevant policy variable is a binary variable indicating whether the individual was granted a Ramón y Cajal contract, that we denote as D_i , which takes on value 1 if the researcher i has been granted a contract and zero otherwise. Our concern is whether the contract status affects the re-

seacher's productivity outcome Y_i in the four-year period after the call. We undertake the analysis using three alternative outcomes variables, measuring scientific performance of researchers. These variables are the number of contributions published in journals listed in the JCR, the average impact of such contributions, and the maximum impact factor among the JCR journals in which she has published.

As it is well known, the ideal evaluation problem, for a given researcher, consists on comparing her two potential outcomes depending on whether she had and she had not a contract, denoted as Y_{0i} and Y_{1i} , respectively. If both counterfactual outcomes were observed for researcher i, the impact of the contract for such researcher would simply be $(Y_{1i} - Y_{0i})$, and we then could calculate the average treatment effect, i.e., the average impact of the contract computing the sample counterpart of $E(Y_{1i} - Y_{0i})$, where E(.) denotes the mean operator (see Rosenbaum and Rubin, 1983).

Since having or not having a contract are mutually exclusive, for each researcher we just observe either $D_i = 1$ or $D_i = 0$, and therefore we just observe her outcome under one of the two situations, i.e.,

$$Y_i = D_i Y_{1i} + (1 - D_1) Y_{0i}. (1)$$

If we consider the regression model

$$Y_i = \alpha + \rho D_i + u_i, \tag{2}$$

where the parameter ρ is the treatment effect, and u_i is an error term capturing unobserved individual differences in scientific productivity. However,

given that for each researcher we just observe one potential outcome, the regression based on observed outcomes would provide the mean differences in outcomes between researchers granted with a contract and researchers without a contract,

$$E(Y_i|D_i=1) - E(Y_i|D_i=0).$$
 (3)

which, unless contract status were purely random, will differ from ρ . We know, though, that contract status depends on researchers' characteristics. In other words, in the regression above, there is a selection bias, so that $E(u_i|D_i) \neq 0$. Essentially, those researchers with a contract are likely to be more productive than researchers without a contract anyway.

As a consequence, if we consider the naive regression (1) to estimate ρ using our sample of applicants by OLS, regressing the observed outcome Y_i on D_i , the resulting naive estimate of ρ will be contaminated by selection bias, so we would exacerbate the impact of the contract. Yet if we have available additional individual information \mathbf{X}_i that determines treatment status, then the potential outcomes might be mean-independent of D_i , conditional on \mathbf{X}_i . We denote it as conditional mean-independent. In such a case, we could consider the augmented specification

$$Y_i = \alpha + \rho D_i + \beta' \mathbf{X}_i + u_i, \tag{4}$$

If conditional on such additional information, average potential outcomes are independent of the treatment status, then $E(u_i|D_i, \mathbf{X}_i) = 0$ holds. Consequently, OLS estimation of (4) will yield a consistent estimate of the impact

of the contract. The information contained in \mathbf{X}_i consists of the observed variables at the time of the call: the researcher curricular information, the time elapsed since the researcher earner her PhD, the quality of her PhD institution, and the researcher's gender. The validity of the regression estimate of the causal effect in (4) relies on absence of unobserved differences between granted and non-granted researchers that affect their potential outcomes.

Alternatively, we can consider matching estimators of the impact of the contract on those researchers who actually obtained a contract, i.e., the average treatment effect on the treated. The idea is to compare the outcomes of granted researchers with the outcomes of selected non-granted researchers that are similar to the first ones except for their contract status. The criterion to match each granted researchers with a non-granted researcher will be the propensity score, i.e., the probability of being granted a contract, conditional on the applicant's characteristics at the time of the call. The validity of our propensity score matching relies on the validity of the Conditional Independence assumption (CIA), i.e., the selection of researchers in the treatment group (granted researchers) or in the control group (non granted researchers) is only based on observables (Rosenbaum and Rubin, 1983). In other words, once we control for these observables (CV quality and other individual characteristics at the time of the call), being granted or not cannot depend on output. Taking into account that our performance measure consists of the CV quality four years after the call, we are quite confident that the CIA holds.

In Table 4, we report OLS linear regression estimates for the full sample of applicants for the three alternative outcomes. For the sake of camparison, we report the naive unconditional estimates of the impact of the contract corresponding to eq. (1), and the conditional estimations of the impact of the contract based on (4). In all estimations we control for differences across scientific areas using a set of area binary dummies. For any of the outcomes, the naive estimates yield a highly positive and significant effect of the contract. It must be noted, though, that there are substantial differences between researchers with and without contract that, on average, make granted researchers more productive than non-granted researchers, so we expect unconditional estimates to be contaminated by a strongly positive selection bias. This is confirmed by the conditional estimates of the impact of the contract, which are much smaller in magnitude. In fact, we find that, when we control for researcher' characteristics at the time of the application, the contract has no significant effect on the number of published contributions in the four-year period after the call, so there are not differences in quantity by contract status. If we regard the influence of the scientific contributions, measured by either the average impact or the maximum inpact, we find that contract status entails significantly positive differences in quality. This results suggest that researchers keep producing scientific output at a similar pace irrespective of their contract status, but researchers with a contract achieve a higher scientific influence.

When we disaggregate by areas, we get similar results than with the full

sample. The naive estimates of the impact of the contract, reported in Table 5, are positive and significant, both for quantity and quality of scientific contributions, in most of the areas. When we condition on researchers' characteristics, there are no longer differences in the quantity of scientific contributions. But some differences across areas arise in the scientific quality. In particular, in the case of Biology, Chemistry and Physics, we find a positive effect of the contract on the average impact and the maximum impact. The effects are not significant for Economics, Engineering and Humanities. In the case of Economics, the sample size is very small. In the case of Engineering and Humanities, there can exist differences in the scientific standards that rule these disciplines. The development of patents is particularly relevant in many Engineering fields, as much as the published contributions in scientific journals.

We have a substantially different result for Medicine, as the contract has a negative effect on the scientific impact, which is significant when we consider maximum impact. This is an intriguing result, which deserves further investigation. It suggests that non-granted researchers keep their research career in an environment that favors their scientific impact. Issues like the availability and quality of laboratories and other available resources can be behind this.

The estimation results for the propensity score matching estimates are shown in Table 7. Granted applicants are then matched with non-granted applicants who are similar in their propensity score. To check whether results might be influenced by the 'similarity' criterion, we consider three alternative criteria: kernel, nearest neighbor and stratification. We have used the Stata procedure written by Becker and Ichino (2002). The variables used to estimate the probability of being granted a contract are the same that we have used as covariates in the conditional regression estimates.⁴ Essentially, alternative matching criteria do not differ except for the significance of the estimated effects: kernel and stratification methods yield very similar results, and nearest neighbor method is typically less precise. Qualitatively, the results resemble the obtained for the conditional regression estimates. Again, the contract status does not affect the quantity of publications, and the contract has a positive impact in our two measures of scientific quality for Biology, Chemistry and Physics areas. Interestingly, contract status for Medicine does not yield significant differences in scientific quality.

In general, the results suggest that, for researchers who are comparable in their ex ante characteristics, the contract status does not yield differences in the number of scientific contributions four years after the call. In addition, researchers with a Ramón y Cajal contract show, on average, a scientific impact four years after the call which higher than for comparable researchers without a contract. This difference is, indeed, significant for several areas. We find, though, an exception in the case of Medicine.

⁴The estimated propensity score yields predicted probabilities of being granted a contract that hold the balancing property. Such property establishes that, conditioning on the propensity score, the distribution of the explanatory variables is not different for granted and not granted researchers.

4 Conclusions

The Ramón y Cajal Program was created to ameliorate the shortage of funds for research personnel and to improve the quality of the Spanish R&D system. For that purpose, the program provided funding to recruit quality researchers and to provide them an entry point into the R&D system. The Program was successful in selecting high quality researchers, the selection being based on curricular merits.

We have analyzed the effect of the program on the productivity of the selected researchers and compare them with scholars with similar curricular characteristics that were not awarded with a Ramon y Cajal contract. We have undertaken two alternative approaches to estimate the causal effect of the contract: conditional regression and propensity score matching procedures. Overall, the results provided by both methods are alike.

We find that the selection process was based on the applicant's research curriculum and that the researchers maintain, once in the Program, a quantitative level of scientific production comparable with similar researchers that were marginally rejected in the selection process. When we consider the scientific impact of the researchers, we find it is at least as high as that for non-granted researchers, and it is significantly higher in several areas.

In some areas for which we do not find a significant effect of the contract, particularly Engineering and Humanities, there can exist different scientific practices and standards. In such cases, our curricular measures can render insufficient to characterize the research merits of the candidates. Also, we find a differential result for Medicine. According to our conditional regression estimates, the contract has a negative causal effect on scientific impact; the matching estimator is also negative but significant. This is an intriguing result; further investigation would require additional data to ascertain where those non-granted researchers that ex ante are comparable with granted researchers have developed their research career in the years following the program call.

Our results point out the success of the Program in increasing the scientific impact of the Spanish system. Interestingly, the program does not appear to have an effect on the quantity of scientific papers produced, but it has on the impact of the scientific contributions. This is an important result, that supports that policies aimed at increasing the stock of human resources in scientific research help to rise the international impact of the Spanish R&D system.

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Table 1. Distribution of applicants and contracts Percentages in each category

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ALL								
$\overline{\text{Applicants}^a}$	6842							
% Contracts ^b	2224							
By Year (%)	2002	2003	2004	2005	2006	2007		
$\overline{\text{Applicants}^a}$	36.6	37.3	19.6	18.7	19.5	21.5		
% Contracts ^b	19.8	27.5	22.1	18.9	18.5	16.4		
By Gender (%)	Fem.	Male						
$\overline{\text{Applicants}^a}$	41.5	58.5						
% Contracts ^b	28.1	36.2						
By Area (%)	Biol.	Econ.	Chem.	Eng.	Medic.	Phys.	Hum.	
$\overline{\text{Applicants}^a}$	35.6	2.4	15.8	11.5	12.2	9.0	13.5	
% Contracts ^b	39.4	31.3	44.8	44.5	33.9	42.2	24.3	
By Ph.D	Up to	3-6	> 6					
tenure (%)	2 yr.	years	years					
$\overline{\text{Applicants}^a}$	15.1	55.6	29.2					
% Contracts ^b	28.2	33.4	32.1					

^aThe percentages of applicants by category add over 100%, since a fraction of them apply in several years and/or in several areas.

 $[^]b\mathrm{Share}$ of granted researchers in the corresponding category.

Table 2. Curricular information of applicants by contract status

	Ave	erage	Ave	erage	Maximum		
	No. of papers		Impac	t factor	Impac	t factor	
CONTRACT:	Yes	no	Yes	no	Yes	no	
All	2.5	1.6	1.8	1.0	2.4	1.2	
	(6.4)	(5.9)	(3.4)	(2.0)	(5.2)	(3.2)	
By Gender							
Female	2.6	1.8	2.1	1.1	2.7	1.4	
	(6.6)	(6.0)	(3.1)	(2.0)	(5.1)	(3.6)	
Male	2.4	1.5	1.7	0.9	2.3	1.1	
	(6.3)	(6.1)	(3.7)	(2.0)	(5.4)	(2.9)	
By Research area							
Biology	2.7	2.5	2.4	1.4	3.1	1.8	
	(6.3)	(8.5)	(4.2)	(2.2)	(5.7)	(4.1)	
Economics	1.9	0.9	0.8	0.5	0.9	0.5	
	(1.9)	(3.6)	(0.9)	(1.4)	(1.0)	(1.5)	
Chemistry	3.0	1.7	1.7	1.0	2.4	1.2	
	(5.2)	(3.9)	(2.5)	(1.5)	(4.1)	(2.3)	
Engineering	0.9	0.5	0.5	0.2	0.4	0.2	
	(2.6)	(1.5)	(1.5)	(0.4)	(1.0)	(0.4)	
Medicine	4.7	2.6	2.7	1.8	4.2	2.1	
	(9.7)	(6.4)	(4.0)	(2.9)	(8.0)	(4.3)	
Physics	2.3	0.8	1.9	0.9	2.1	1.0	
	(9.6)	(2.1)	(3.0)	(2.3)	(3.7)	(2.5)	
Humanities	0.5	0.2	0.8	0.2	1.2	0.2	
	(1.4)	(0.9)	(2.9)	(0.6)	(6.0)	(0.7)	
By Ph.D. tenure							
Up to 2 years	0.9	0.6	1.2	0.6	1.3	0.6	
	(2.2)	(2.3)	(3.2)	(1.7)	(3.5)	(2.2)	
3-6 years	2.1	1.4	2.0	1.0	2.7	1.3	
	(4.1)	(5.0)	(3.4)	(2.1)	(5.5)	(3.2)	
More than 6 years	3.6	2.7	1.9	1.3	2.7	1.6	
	(9.6)	(8.7)	(3.6)	(2.2)	(5.7)	(4.0)	

Standard deviation in parentheses below the sample mean of each category.

Table 3. Assessment of candidates

Dependent variable: Score By Area $ALL^{(1)}$ Econ. Phys. Biol. Chem. Eng. Hum. Medic. 58.47^{\S} 49.68^{\S} 56.72^{\S} 46.29^{\S} 68.25^{\S} 53.82^{\S} Constant 53.83^{\S} 54.92^{\S} (1.56)(2.28)(7.52)(3.37)(3.67)(3.88)(3.86)(4.36)0.02 -0.011.20 0.06 1.06^{\dagger} 0.11^{*} 1.21 # Papers -0.04(0.05)(0.04)(1.30)(0.21)(0.52)(0.12)(0.06)(1.45) 0.68^{\S} 0.92^{\S} -3.06 1.69^{\S} -1.84*0.90Avg. IF 0.220.15(0.21)(0.31)(7.40)(0.63)(0.97)(0.22)(1.44)(1.96) 5.75^{\S} 0.33^{\dagger} Max. IF 0.17-0.142.560.230.620.06(0.11)(0.18)(5.87)(0.34)(2.08)(1.00)(0.69)(0.14) 3.54^{\S} 4.23^{\S} Gender -0.93 5.06^{\S} 5.20^{\dagger} 2.96^{\dagger} -0.09 4.34^{\dagger} (0.60)(0.89)(5.19)(1.85)(2.55)(1.28)(2.13)(1.89) 1.46^{\S} 1.57^{\S} -1.71 2.40^{\S} 1.58^{\S} 1.36^{\dagger} 2.33^{\S} PhD tenure 0.86*(0.18)(0.34)(1.40)(0.44)(0.50)(0.60)(0.54)(0.40)PhD tenure² -0.05^{\S} -0.05^{\dagger} 0.10^{*} -0.08^{\S} -0.03 -0.09^{\S} -0.06^{\dagger} -0.07^{\S} (0.01)(0.02)(0.05)(0.02)(0.02)(0.03)(0.03)(0.01) 0.54^{\S} 0.61^{\S} 0.66^{\S} 0.64 0.53^{\dagger} 0.52*0.130.10PhD center (0.17)(0.41)(0.26)(0.22)(0.29)(0.29)(0.24)quality (0.09)# obs. 1,995 394 730 616 4,967 104 694 434 R^2 0.940.940.920.92 0.96 0.950.91 0.89

Binary dummies for esearch areas (DGI) included in the full sample regression.

Gender takes on value 1 if male and 0 otherwise. Robust standard error in parentheses.

^{*,&}lt;sup>†</sup>, Significant at 10%, 5% and 1% levels, respectively.

Table 4. Applicant's scientific output 4 years after the call. OLS estimates. Full sample

Outcome variable	No. of	papers	Avg. Impact		Max.	Impact
	(i)	(ii)	(iii)	(iv)	(v)	(vi)
Contract	0.785^{\S}	0.050	0.637^{\S}	0.223^{\S}	0.849^{\S}	0.286^{\dagger}
	(0.225)	(0.140)	(0.096)	(0.079)	(0.134)	(0.121)
No. of papers (t-4)		0.940^{\S}		0.036^{\S}		0.066^{\S}
		(0.048)		(0.011)		(0.019)
Average IF (t-4)		-0.073		0.393^{\S}		0.449^{\S}
		(0.086)		(0.094)		(0.124)
Maximum IF (t-4)		0.042		0.058*		0.091
		(0.069)		(0.030)		(0.058)
Gender		0.206*		-0.040		0.022
		(0.117)		(0.071)		(0.103) .
PhD tenure		-0.241^{\S}		-0.048^{\S}		-0.071^{\S}
		(0.033)		(0.016)		(0.026)
PhD tenure ²		0.006^{\S}		-0.002^{\dagger}		0.002*
		(0.001)		(0.001)		(0.001)
PhD center quality		-0.028^{\dagger}		-0.004		-0.002
		(0.013)		(0.009)		(0.014)
Number of observations		3, 165				
R^2	0.09	0.80	0.22	0.51	0.19	0.44

Dummies for research areas and residence included in multiple regressions.

Gender takes on value 1 if male and 0 otherwise. Robust standard error in parentheses.

^{*,} † , Significant at 10%, 5% and 1% levels, respectively.

Table 5. Applicant's scientific output 4 years after the call. Unconditional OLS estimates. By research areas.

Outcome variable: Number of papers									
	Biol.	Econ.	Chem.	Eng.	Medic.	Phys.	Hum.		
Contract	0.039^{\S}	1.005	1.188	0.463^{\dagger}	2.376^{\dagger}	1.285*	0.113		
	(0.402)	(0.650)	(0.494)	(0.230)	(1.125)	(0.676)	(0.216)		
R^2	0.09	0.17	0.15	0.15	0.15	0.09	0.06		
Outcome	Outcome variable: Average Impact								
Contract	0.689^{\S}	0.116	0.544^{\dagger}	0.195^{\S}	0.614*	1.378^{\S}	0.195		
	(0.182)	(0.204)	(0.151)	(0.063)	(0.319)	(0.494)	(0.142)		
R^2	0.28	0.26	0.37	0.23	0.34	0.16	0.08		
Outcome	variable:	Maximu	m Impact						
Contract	0.960^{\S}	0179	0.826^{\S}	0.263^{\S}	0.705	1.686^{\S}	0.332		
	(0.258)	(0.279)	(0.210)	(0.079)	(0.480)	(0.603)	(0.224)		
R^2	0.24	0.22	0.33	0.22	0.26	0.16	0.07		
Obs.	1,308	90	552	416	448	329	494		

Robust standard errors in parentheses.

^{*,†,} \$Significant at 10%, 5% and 1% levels, respectively.

Table 6. Applicant's scientific output 4 years after the call. Conditional OLS estimates. By research areas.

Outcome variable: Number of papers									
	Biol.	Econ.	Chem.	Eng.	Medic.	Phys.	Hum.		
Contract	-0.042	-0.024	0.107	0.167	-0.115	0.241	-0.231		
	(0.229)	(0.596)	(0.431)	(0.234)	(0.518)	(0.252)	(0.156)		
R^2	0.85	0.86	0.63	0.50	0.85	0.88	0.46		
Outcome	Outcome variable: Average Impact								
Contract	0.300^{\dagger}	-0.240	0.314^{\dagger}	0.099	-0.130	1.122^{\dagger}	-0.037		
	(0.145)	(0.244)	(0.146)	(0.076)	(0.223)	(0.495)	(0.089)		
R^2	0.52	0.69	0.59	0.47	0.68	0.38	0.40		
Outcome	variable:	Maximur	n Impact						
Contract	0.501^{\dagger}	-0.313	0.528^{\dagger}	0.154	-0.705^*	1.218^{\dagger}	-0.065		
	(0.232)	(0.390)	(0.217)	(0.096)	(0.379)	(0.537)	(0.111)		
R^2	0.44	0.57	0.52	0.66	0.56	0.35	0.49		
Obs.	1, 189	81	464	304	417	269	441		

Covariates include dummies for research areas and residence, gender and PhD tenure, as well as curricular variables at the time of the call. Robust standard errors in parentheses.

^{*,&}lt;sup>†</sup>, Significant at 10%, 5% and 1% levels, respectively.

Table 7. ATT Estimates of the impact of the contract Propensity score matching estimates

Matching	Outcome variable: Number of papers									
Method	All	Biol.	Econ.	Chem.	Eng.	Medic.	Phys.	Hum.		
kernel	0.168	0.232	-0.283	-0.024	0.479	0.673	1.075	-0.112		
	(0.160)	(0.223)	(3.009)	(0.786)	(0.275)	(1.334)	(0.725)	(0.365)		
NN	0.291	0.283	0.143	0.310	0.217	0.026	1.188	-0.237		
	(0.489)	(0.609)	(3.642)	(1.186)	(0.380)	(2.002)	(0.849)	(0.503)		
stratif.	0.031	0.277	1.320	0.248	0.316	0.919	0.855	-0.103		
	(0.346)	(0.343)	(0.529)	(0.845)	(0.235)	(1.645)	(0.779)	(0.224)		
	Outcom	e variable	: Average	e Impact						
kernel	0.382^{\S}	0.499^{\S}	-0.328	0.320^{\dagger}	0.086	0.188	1.254^{\dagger}	0.154		
	(0.096)	(0.183)	(0.740)	(0.148)	(0.101)	(0.310)	(0.514)	(0.140)		
NN	0.306^{\dagger}	0.492^{\dagger}	-0.447	0.291	-0.041	-0.140	0.920	0.172		
	(0.141)	(0.247)	(1.033)	(0.213)	(0.142)	(0.414)	(0.625)	(0.173)		
stratif.	0.351^{\S}	0.489^{\dagger}	0.057	0.279^{*}	0.111^*	0.232	0.892^{\dagger}	0.054		
	(0.115)	(0.214)	(0.236)	(0.155)	(0.069)	(0.349)	(0.445)	(0.121)		
	Outcom	e variable	: Maximi	ım Impac	t					
kernel	0.494^{\S}	0.774^{\S}	-0.376	0.552^{\dagger}	0.156	-0.270	1.262^{\dagger}	0.277		
	(0.133)	(0.258)	(1.199)	(0.224)	(0.118)	(0.549)	(0.522)	(0.235)		
NN	0.448^{\dagger}	0.845^{\dagger}	-0.750	0.476	0.021	-0.506	0.995	0.294		
	(0.209)	(0.367)	(1.582)	(0.298)	(0.172)	(0.838)	(0.714)	(0.239)		
stratif.	0.458^{\S}	0.742^{\dagger}	0.030	0.516^{\dagger}	0.171^{*}	-0.200	0.880*	0.096		
	(0.156)	(0.311)	(0.393)	(0.211)	(0.096)	(0.570)	(0.546)	(0.151)		

Propensity score covariates include curricular variables at the time of the call, gender, PhD tenure, and quality of the PhD center.

The three alternative matching criteria are kernel, NN (nearest neighbor) and stratification. Bootstrap standard errors, based on 500 replications, in parentheses.

 $^{^*}$, † , Significant at 10%, 5% and 1% levels, respectively.

Table A1

Research areas (DGI)

Physics and Space Sciences

Earth Sciences

Materials Science and Technology

Chemistry

Chemical Technology

Plant and Animal Biology. Ecology

Agriculture

Livestock and Fishery

Food Science and Technology

Molecular and Cell Biology and Genetics

Physiology and Pharmacology

Medicine

Mechanical, Ship and Aeronautical Engineering

Electrical and Electronic Eng. and Robotics

Civil Engineering and Architecture

Mathematics

Computer Sciences

Information and Communication Technologies

Economics

Law

Social Sciences

Psychology and Education Sciences

Philology and Philosophy

History and Art

Table A2

Research areas (Publish or Perish)

Biology, Life Sciences, Environmental Science Business, Administration, Finance, Economics Chemistry and Materials Science Engineering, Computer Science, Mathematics Medicine, Pharmacology, Veterinary Science Physics, Astronomy, Planetary Science

G : 1 G : A + TT :::

Social Sciences, Arts, Humanities