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# On heuristic solutions for the stochastic flowshop scheduling problem* 

Jose M. Framinan, Paz Perez-Gonzalez

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#### Abstract

We address the problem of scheduling jobs in a permutation flowshop when their processing times adopt a given distribution (stochastic flowshop scheduling problem) with the objective of minimisation of the expected makespan. For this problem, optimal solutions exist only for very specific cases. Consequently, some heuristics have been proposed, all of them with similar performance. In our paper, we first focus on the critical issue of estimating the expected makespan of a sequence and found that, for instances with a medium/large variability (expressed as the coefficient of variation of the processing times of the jobs), the number of samples or simulation runs used in the literature may not be sufficient to derive robust conclusions. We thus propose a procedure with a variable number of iterations that ensures that the percentual error in the estimation of the expected makespan is bounded with a very high probability. Using this procedure, we test the main heuristics proposed in the literature and find significant differences in their performance, in contrast with existing studies. We also find that the deterministic counterpart of the most efficient heuristic for the stochastic problem performs extremely well for most settings, which indicates that (at least within the limitations of our study), a practical way to solve the stochastic problem may be to simplify it to its deterministic version.


Keywords: Scheduling, Flowshop, Stochastic, Makespan Objective, Heuristics

[^0]
## 1 Introduction

The flowshop scheduling problem with makespan objective (usually denoted as $\operatorname{Fm} \mid$ prmu $\mid C_{\max }$ ) has been subject of research for more than 60 years, being one of the most comprehensively studied problems in Operations Research (see in this regard the reviews by Framinan et al., 2004, Reza Hejazi and Saghafian, 2005 and Ruiz and Maroto, 2005). This decision problem consists of how to schedule jobs in a permutation flowshop in order to minimize the maximum completion time or makespan. A classical assumption is that the processing times of each job in each machine are considered different, but known in advance (deterministic). In contrast, our paper deals with the problem of scheduling $n$ jobs in a permutation flowshop consisting of $m$ machines where the processing times are not deterministic, but follow some known distribution. The objective considered is that of minimizing the expected makespan. This problem is considered to be more realistic that their deterministic counterpart, as it allows capturing part of the inherent variability present in many real-life manufacturing environments (see e.g. Hopp and Spearman, 2008). In the following, we will denote our problem as $F m|p r m u| E\left[C_{m a x}\right]$.

The $F m|p r m u| E\left[C_{m a x}\right]$ problem has been much less studied than its deterministic counterpart, and it is clearly much more complex. In fact, apart from a dominance rule obtained by Makino (1965) for the case of two jobs, no exact solution is available without assumptions on the distribution of the processing times. For $m=2$ and exponential distribution of the processing times, Talwar (1967) conjectured an exact solution for the problem that was later proved to be optimal by Cunningham and Dutta (1973), and is currently known as Talwar's rule.

Despite these advances, for the rest of the cases, no optimal procedure has been found. For the two-machine case, three approximate solutions have been proposed by Baker and Trietsch (2011) based both in Talwar's rule and in Johnson's rule (Johnson, 1954) for the deterministic flowshop, all of them with similar (near optimal) performance. For the general machine case, Baker and Altheimer (2012) suggest three heuristics based on adaptations of the CDS (Campbell et al., 1970) and NEH (Nawaz et al., 1983) heuristics, again with similar and near optimal performance. Although it might seem that, from these results, the problem Fm|prmu|E[C max $]$ is already solved, some issues have to be discussed:

- First of all, the evaluation of sequences in an stochastic flowshop is far from being a trivial task. Since the objective is to obtain the expected makespan of a given sequence, $E\left[C_{\text {max }}\right]$ has to be estimated by running $N$ simulations using the sequence as a solution, from which a sample $C_{\text {max }}^{i}(i=1, \ldots, N)$ is obtained. Then, $E\left[C_{\max }\right]$ is estimated by averaging the sample, i.e. $E\left[C_{\max }\right] \approx \bar{C}_{\max }=\frac{1}{N} \sum_{i=1}^{N} C_{\max }^{i}$.

Up to now, there is no standardised procedure to determine $N$, although the authors of related contributions use a large number in order to ensure the significance of the estimation. Thus, Baker and Altheimer (2012) use $N=100,000$ whereas Gourgand et al. (2003) employ $N=200,000$ regardless the size and characteristics of each instance, while Portougal and Trietsch (2006) set $N$ to 10,000 for the 2 -machine case. In addition, there is no mechanism to establish the statistical significance of the so-obtained $\bar{C}_{\text {max }}$ and, consequently, to assess the differences in the performance among different heuristics.

- Due to the computational complexity of the stochastic problem, the experiments in the literature have been limited to very small problem sizes (up to $n=10$ and $m=6$ in the most recent studies). This makes the conclusions obtained so far to be restricted to very small problem sizes, and perhaps not valid for bigger problem sizes.
- Finally, to the best of our knowledge, the neccessity of heuristics specifically designed for the stochastic problem has not been yet determined. In other words, one may try to solve the stochastic flowshop scheduling problem by transforming it into its deterministic counterpart, i.e. by obtaining a flowshop with the same number of jobs and machines but with deterministic processing times equal e.g. to the means of those from the stochastic problem. Then, heuristics for the $F m|p r m u| C_{\max }$ problem can be applied and a (possibly good) sequence for the deterministic problem can be obtained. If this sequence performs well when applied to the stochastic problem, then the need of specific stochastic heuristics can be questioned. However, such test has not been conducted so far. It is foreseable that the so-obtained sequences perform worse that those specifically designed for the stochastic
version, but maybe the differences in the quality of the results do not justify the much higher computation times required for the stochastic heuristics. Even if the deterministic heuristics are not valid for some cases, it would be interesting to quantify the degree of variability for which using them is still acceptable, as it is clear that an stochastic flowshop with low variability would resemble very much to a deterministic flowshop.

With these issues in mind, we first discuss and propose a procedure for estimating the expected makespan of a sequence in an stochastic flowshop, so the error in such estimation is bounded by a given percentage. In this way, the statistical significance of the results obtained by the different heuristics can be more clearly established. Interestingly, the results show that the sample sizes ( $N$ ) obtained from our procedure proposed range from very small to very high values, thus supporting the conclusion that no predetermined value can be easily found regardless the variability of the instances and the error assumed in the estimation.

Next, we compare the main heuristics proposed in the literature as well as the expected makespan obtained from the application of purely deterministic procedures to the mean processing times of the stochastic problem. These heuristics are tested for problem sizes larger than those presented so far in the literature, so the conclusions from the results can be better supported. The results show that, in contrast to Baker and Altheimer (2012), there are significant differences in the performance of the heuristics, and that -perhaps not so surprisingly- the performance of the sequences obtained from purely deterministic methods in the stochastic flowshop do not differ greatly from that obtained from specific stochastic methods.

The remainder of the paper is organised as follows: First, the problem under consideration is formally described in Section 2, where the main contributions of the literature are discussed. Since our work is of computational nature, we devote Section 3 to discuss the key issue of the testbed in which the heuristics are to be compared, as we intend to capture different problem sizes and different degrees of variability of the flowshop. Next, we present in Section 4 the procedure to estimate the expected mean makespan of a given solution, and compare the number of iterations required with those employed in the literature. The comparison of the performance of different heuristics for the problem is done in Section 5, where the main results are also discussed. Finally,

Section 6 present the conclusions and points out future research lines.

## 2 Background

A flowshop consists of $n$ jobs that must be processed on $m$ machines in the same order, where job $i$ requires $p_{i j}$ time units to be processed on machine $j$. The scheduling problem in flow shops is to find a sequence of jobs for each machine according to certain performance measure(s). Additionally, for many situations, it is assumed that the job sequences will be the same on every machine (permutation flowshops). Other hypotheses common in scheduling research are, e.g. the simultaneous availability of all jobs and of all machines, deterministic processing times, etc. For a complete list of these assumptions, see e.g. Framinan et al. (2004).

While the deterministic flowshop scheduling problem with makespan objective has been extensively studied (see the reviews mentioned above), the same cannot be said about its stochastic counterpart. For the two-jobs case and a general distribution of the processing times, a dominance rule is given by Makino (1965), but this result is extremely restrictive and with little applicability for most practical settings.

By making assumptions on the distribution of the processing times of the jobs, an important result is due to Talwar (1967). He conjectures that the expected makespan is minimized when the processing time of the jobs follows an exponential distribution by sequencing the jobs in non increasing order of $\frac{1}{\mu_{i 1}}-\frac{1}{\mu_{i 2}}$, where $\mu_{i j}$ is the mean processing times of job $i$ on machine $j$. This order is proved to be optimal by Cunningham and Dutta (1973), and it is currently known as Talwar's rule. As an extension of this rule to other distributions, Kalczynski and Kamburowski (2006) heuristically adapt Talwar's rule for the Weibull distribution. For a general family of distributions, Portougal and Trietsch (2006) develop a heuristic named PSH which starts with the solution given by Johnson's rule (Johnson, 1954) for the deterministic flowshop, and applies an adjacent pairwise interchange (a reason why this heuristic is later renamed API by Baker and Trietsch, 2011). Finally, Baker and Trietsch (2011) test three different procedures for different families of distributions, i.e.: Talwar's rule, Johnson's rule, and the API heuristic. They conclude that the (deterministic) Johnson's rule could be better unless the coefficient of variation
of the jobs is very high. For such cases, Talwar's rule or API may perform better.
For the general flowshop problem with $m$ machines, Baker and Altheimer (2012) propose different heuristics. The first heuristic is called CDS/Johnson and consists of obtaining a set of m-1 2-machine flowshop subproblems with the addition of the processing times of the jobs in the manner of the CDS heuristic by Campbell et al. (1970). More specifically, 2-machine flowshop subproblem $k$ (with $k=1, \ldots, m-1$ ) is constructed by obtaining the processing times of job $i$ in the first (second) machine of this subproblem as $A_{i}=\sum_{j=1}^{j=k} p_{i j}\left(B_{i}=\sum_{j=k+1}^{j=m} p_{i j}\right)$. Then, Johnson's procedure is applied to each of the resulting $m-1$ subproblems, and $m-1$ sequences are obtained. The estimation of the expected makespan of each of this sequences is obtained (in their case by running 100,000 simulations and taking the average makespan value), and the one yielding the lowest value is selected.

The second tested heuristic is the CDS/Talwar heuristic. In a similar manner to the previous one, a set of $m-12$-machine flowshop subproblems are obtained, and a sequence is obtained for each one by applying Talwar's rule. Out of these $m-1$ sequences, the one with the minimum estimation of the expected makespan is selected.

The third heuristic tested is based on the famous NEH heuristic proposed by Nawaz et al. (1983) for the deterministic case. This heuristic consists of two phases: First the jobs are ranked according to the descending sum of their mean processing times. In a second phase, a solution is constructed as follows: Starting from a partial sequence constructed by taking the first job of the rank, then, for $k=2, \ldots, n, k$ partial sequences are constructed by inserting the $k$-th job of the rank in all $k$ slots of the partial sequence. These $k$ partial sequences are evaluated with respect to their expected makespan (estimated, as in previous cases, by means of obtaining a sample via simulation and taking the average value), and the one obtaining the lowest value of the estimation of the expected makespan is retained as partial sequence for step $k+1$. The procedure is repeated until a full sequence is obtained.

Among the three heuristics, Baker and Altheimer (2012) find that the NEH adaptation is the best one, but none of the heuristic procedures dominates the others. Furthermore, their performance is compared against that of a genetic algorithm (assumed to find the optimal or near-optimal sequence for most instances), and the authors conclude that the three heuristics
generate average suboptimalities of less than $1 \%$, leaving little room for the development of new approximate algorithms.

Despite the advances reported, there are some issues already discussed in Section 1 that affect the existing results and deserve further research. A closer look on how to estimate the expected makespan is needed, in order to add statistical consistency to the results. Furthermore, the neccessity of special stochastic heuristics for the problem has not been established, particularly when, for the two-machine case, Portougal and Trietsch (2006) state the excellent performance of the Johnson's deterministic heuristic in the stochastic setting. To address all these issues, we first need a benchmark testbed to conduct the computational experience. The design of this testbed is presented in the next section.

## 3 Testbed design

Regarding the design of a testbed for flowshop scheduling, we first have to decide about the problem sizes, i.e. the number of jobs $n$ and machines $m$ of the different instances. Given the computational complexity of the stochastic model, problem sizes have to be much more reduced than those in the deterministic counterpart, however we want to subtantially increase the existing problem sizes in order to gain generality on the results. With these premises, we choose $n \in\{5,10,15,20\}$ and $m \in\{2,5,10,20\}$.

In order to ease the analysis, we assume that all distribution of the processing times belong to the same family, which is a common assumption almost universally made (see e.g. Gourgand et al., 2003, Baker and Trietsch, 2011 or Baker and Altheimer, 2012). Regarding to the family of distributions, there are several option, including the random distribution (Baker and Trietsch, 2011 or Baker and Altheimer, 2012), the normal distribution (Gourgand et al., 2003) the exponential distribution (Gourgand et al., 2003, Baker and Trietsch, 2011 or Baker and Altheimer, 2012), and the lognormal distribution (Baker and Trietsch, 2011 or Baker and Altheimer, 2012).

After reviewing the different distributions, we will assume a $\log$ normal distribution for our testbed. The log normal distribution is characterised by two parameters (mean $\mu$ and standard deviation $\sigma$ ), and it has a considerable practical value for modelling real-life processing times.

Additionally, in contrast e.g. to the exponential distribution, we can control its variability by means of the standard deviation and thus model different degrees of variability of the instances in the testbed.

Regarding the average processing times of the jobs $\left(\mu_{i j}\right)$, each one is drawn from a discrete uniform $[1,99]$ distribution. In the deterministic counterpart, these values are assumed to generate difficult instances (see e.g. Dannenbring, 1977, or Campbell et al., 1970), so we expect the same for the stochastic case. For each mean processing time, we want to model different degrees of variability. Setting different values for the standard deviation in an straightforward manner could be rather tricky and produce non realistic processing times. Therefore, we use different values of the coefficient of variation $c=\frac{\sigma}{\mu}$ to generate several instances. More specifically, $c \in\{0.01,0.1,0.2,0.5,1\}$ to capture the different scenarios of variability. When $c=0.01$, the processing times are close to be deterministic, but for $c=1$ the processing times for a job suffer a great variation. Although bigger $c$ values are naturally possible, we believe that these do not represent reasonably realistic environments, as very large coefficient of variations are not the norm in industry due to the substantial efforts done to reduce the variability of the processing times via process automation and standardisation procedures. Note that the case $c=1.0$ yields the same variance as in the exponential distribution, and, in order to make sure that we also cover this case, in Section 4 we also include results assuming an exponential distribution of the processing times.

In summary, we build each instance of the testbed in the following manner: For a given coefficient of variation $c$, we select a number of jobs and a number of machines (problem size). Next, for each job on each machine, we generate their mean processing time $\mu_{i j}$ by using an uniform $[1,99]$ distribution. The standard deviation of each job on each machine $\sigma_{i j}$ is then set to $c \cdot \mu_{i j}$. This procedure is repeated in order to generate 20 instances for each problem size and coefficient of variation. In total, 1,600 instances (320 for each value of $c$ ) are generated.

## 4 A procedure for the estimation of the expected makespan

As mentioned before, a critical issue in stochastic scheduling is how to evaluate the solutions. Recall that here the objective function is the expected makespan, therefore, given a sequence, an expected makespan must be assigned to this sequence. However, such expected makespan has to be estimated via a sample mean of the makespans corresponding to this sequence. The usual way to obtain the sample mean is to conduct a very large number of simulations of the makespan. Gourgand et al. (2003) assesed the accuracy of such simulations by making comparisons of the simulation results against those obtained by a Markov chain, and found that sample sizes of 200,000 produced $95 \%$ confidence intervals of the order of $0.1 \%$. In their experiments, Baker and Altheimer (2012) use a sample size of 100,000 for problem sizes where $m \in\{2,3,6\}$ and $n$ is up to 10 jobs whereas, for the 2-machine case, Portougal and Trietsch (2006) use a sample size of 10,000 . Nevertheless, it is clear that the sample size should depend -at least- on two factors, namely the confidence interval of the results, and the stochasticity of the problem. Note that the way the solutions are evaluated may determine the significance of the results, here the criticality of this issue.

In this paper, we propose a method to estimate the expected makespan of a sequence based on the maximum percentual error accepted for the estimation of $E\left[C_{\max }\right]$. More specifically, the halfwidth of a confidence interval for $E\left[C_{\max }\right]$ of $1-\alpha$ confidence level is given by $t_{\alpha / 2, N-1} \frac{s}{\sqrt{N}}$, where $s$ is the sample standard deviation of the makespan, and $\alpha / 2$ is the area of a Student's t-distribution with $N-1$ degrees of freedom left in the interval ( $-\infty, t_{\alpha / 2, N-1}$ ] (see e.g. Vijay and Saleh, 2011). We intend that $t_{N-1, \alpha / 2} \frac{s}{\sqrt{N}} \leq \bar{C}_{\text {max }} \cdot p$, where $p$ is a (small) percentage. By doing so, $E\left[C_{\text {max }}\right]$ is confined in the interval $\left[\bar{C}_{\max }(1-p), \bar{C}_{\max }(1+p)\right]$ with a $1-\alpha$ confidence level. If we set $\alpha$ to a very low value (in our experiments, $\alpha=0.001$ ), we can be almost sure (statistically speaking) that $p$ represents the maximum relative error of the estimation of $E\left[C_{\max }\right]$ and thus use $p$ to check the significance of different results obtained by several heuristics. Note that the normality assumption of the sum of the sample values of makespan required to use this confidence interval seems a very loose restriction given the high number of samples (simulations) that would be run in practice, and the fact that the result of each run is independent from the others.

More specifically, our procedure for estimating the expected makespan of a sequence $S$ consists of the following steps:

1. Set the simulations counter to zero, i.e. $N:=0$
2. Set the sum of makespans to zero, i.e. $S u m C_{\max }:=0$
3. Set the sum of squares of makespans to zero, i.e. $S u m S C_{\max }:=0$
4. do:
(a) Run a simulation to obtain a sample makespan $C_{\max }$ of $S$.
(b) Update the number of simulations, i.e. $N:=N+1$.
(c) Update the sum of makespans, i.e. $S u m C_{\max }:=S u m C_{\max }+C_{\max }$
(d) Update the sum of squares of makespans, i.e. $S u m S C_{\max }:=S u m S C_{\max }+C_{\max }^{2}$
(e) Calculate $\bar{C}_{\max }:=\frac{S u m C_{\max }}{N}$ and $s:=\sqrt{\frac{S u m S C_{\max }-N \cdot\left(\bar{C}_{\max }\right)^{2}}{N-1}}$.
while $\frac{s \cdot t_{N-1}, \alpha / 2}{\bar{C}_{\text {max }} \sqrt{N}}>p$
5. Return $\bar{C}_{\text {max }}$.

Note that, technically, we have to force the procedure to run more than one simulation. From the second simulation on, the percentual error oscillates until it is below the desired quantity $p$. Note also that setting unrealistic values of $p$ and $\alpha$ may cause the loop to enter into a deadlock. We can be there sure (with a confidence level of $1-\alpha$ ) that the percentage error in the estimation of the expected makespan is below $p$. In our experiments, $\alpha=0.001$, so that means that we can be quite confident $(99,9 \%)$ on the bounds of the error in the estimation.

In order to check the number of simulation runs required by this procedure for different degrees of variability and percentage error $p$, we obtain the estimations of the expected makespan of a random sequence for each instance in the testbed. The results are shown in Table 1.

From the results, it can be seen that, as foreseable, the number of runs varies greatly depending on the percentage error accepted. Allowing a $5 \%$ error means that results can be obtained with

```
\(N \longleftarrow 0 ; \backslash \%\) Set the simulation counter to zero
\(\pi \longleftarrow \pi_{1}^{\prime}\);
for \(k=2\) to \(n\) do
    \(r \longleftarrow \pi_{k}{ }^{\prime} ;\)
    Determine the values of \(e_{i j}, q_{i j}\) and \(f_{i l}\) from Taillard's acceleration (see equations ??,
    ??, and ??);
    Determine minimal makespan resulting from inserting job \(r\) in all possible positions of
    \(\pi\);
    \(b p \longleftarrow\) First position where the makespan is minimal;
    \(t b \longleftarrow\) Number of positions with minimal makespan (i.e. number of ties);
    \(p t b \longleftarrow\) Array (of length \(t b\) ) with the positions where the makespan is minimal;
    \(i t_{b p}\) is the idletime corresponding to the \(b p\) and set to a very large number;
    if \(t b>1\) and \(k<n\) then
        for \(l=1\) to \(t b\) do
            \(i t^{\prime \prime} \longleftarrow 0 ;\)
            if \(p t b[l]=k\) then
                for \(i=2\) to \(m\) do
                    \(i t^{\prime \prime} \longleftarrow i t^{\prime \prime}+f_{i, k}-e_{i, k-1}-t_{i, r} ;\)
                end
            else
                \(f_{1, p t b[l]}^{\prime} \longleftarrow f_{1, p t b[l]}+p_{1, p t b[l]} ;\)
                for \(i=2\) to \(m\) do
                \(i t^{\prime \prime} \longleftarrow i t^{\prime \prime}+f_{i, p t b[l]}-e_{i, p t b[l]}+p_{i, p t b[l]}-t_{i, r}+\max \left\{0, f_{i-1, p t b[l]}^{\prime}-f_{i, p t b[l]}\right\} ;\)
                \(f_{i, p t b[l]}^{\prime} \longleftarrow \max \left\{f_{i-1, p t b[l]}^{\prime}, f_{i, p t b[l]}\right\}+p_{i, p t b[l]} ;\)
            end
        end
        if \(i t_{b p}>i t^{\prime \prime}\) then
                \(b p \longleftarrow p t b[l] ;\)
                \(i t_{b p} \longleftarrow i t^{\prime \prime} ;\)
        end
        end
    end
    \(\pi \longleftarrow\) Array obtained by inserting job \(r\) in position \(b p\) of \(\pi\);
end
```

Figure 1: Our Tie-Breaking Method

|  |  | $=0.005$, several $c$ v |  |  |  | $=0.01$, several $c$ v |  |  |  | $=0.05$, several $c$ values |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $m$ | 01 | 0.1 | 0.2 | 0.5 | 01 | 0.1 | 0.2 | 0.5 | 0.01 | 0. | 0.2 | 0.5 |
| 5 | 2 | 1975 | 5822 | 753 | 2023060 | 07999 | 461 | 4159 | 0029 | 4327 | 465 | 4981 | 770 |
|  |  | 432423 | 05 | 612 | 135706 | 11 | 101 | 117412 | 324 | 332 | 4413 | 4707 | 420 |
|  | 10 | 432690 | 376 | 45 | 1013 | 08 | 109431 | 113899 | 247341 | 4334 | 4385 | 4559 | 38 |
|  | 20 | 432 | 435 | 445717 | 787848 | 08 | 10891 | 111445 | 9720 | 4336 | 4364 | 446 | 18 |
|  | 2 | 432495 | 402 | 46907 | 1387 | 108129 | 11006 | 117269 | 33399 | 43 | 41 | 47 | 12204 |
| 10 | 5 | 432679 | 437646 | 455 | , | 0817 | 09 | 139 | 448 | 43 | 438 | 456 | 358 |
|  | 10 | 432818 | 612 | 447910 | 804226 | 08210 | 109040 | 111982 | 201065 | 4336 | 析 | 44 | 671 |
| 10 | 20 | 432 | 434916 |  | 69190 | 08233 | 08 | 11056 | 17 | 43 | 4357 | 44 | 6724 |
| 15 |  | 432686 | 73 |  | 105215 | 08177 | 109 | 114367 | 263858 | 433 | 4387 | 4579 | 5 |
| 15 |  | 432793 | 36369 | 449396 | 70760 | 8204 | 109097 | 112343 | 215040 | 4335 | 4371 | 4501 | 7987 |
| 15 | 10 | 432870 | 435250 | 120 | 725289 | 8223 | 108816 | 031 | 81546 | 336 | 360 | 4447 | 498 |
|  | 20 | 432949 | 434 | 440394 | 13 | 08243 | 086 | 1010 | 58538 | 43 | 4353 | 441 | 360 |
|  | 2 | 432788 | 43675 | 51 | 4973 | 08202 | 0919 | 12862 | 23967 | 43 | 37 | 45 | 842 |
| 20 | 5 | 432856 | 43562 | 445 | 765 | 08219 | 1089 | 1146 | 9436 | 43 | 36 | 446 | 177 |
| 20 | 10 | 43 | 4 |  |  | 4 | 108 | 532 | 17275 | 43 | 356 | 4430 | 8 |
| 20 | 20 | 432968 | 434284 | 439143 | 61 | 108248 | 108577 | 109790 | 154052 | 4337 | 4350 | 4399 | 6195 |
|  |  | 432729 | 43714 | 453276 | 9616 | 108188 | 10929 | 113322 | 237649 | 433 | 4379 | 4540 | 9023 |

Table 1: Estimation of the expected makespan for the lognormal testbed: Number of simulation runs required

|  |  | $\mathrm{p}=0.005$ | $\mathrm{p}=0.01$ | $\mathrm{p}=0.05$ |
| ---: | ---: | ---: | ---: | ---: |
| 5 | 2 | 499372 | 124843 | 5016 |
| 5 | 5 | 466236 | 116973 | 4667 |
| 5 | 10 | 452409 | 113241 | 4533 |
| 5 | 20 | 443656 | 111020 | 4443 |
| 10 | 2 | 468839 | 116955 | 4696 |
| 10 | 5 | 452568 | 113256 | 4535 |
| 10 | 10 | 444945 | 111386 | 4457 |
| 10 | 20 | 440093 | 110088 | 4408 |
| 15 | 2 | 457536 | 114205 | 4591 |
| 15 | 5 | 446817 | 111795 | 4478 |
| 15 | 10 | 441605 | 110480 | 4424 |
| 15 | 20 | 438473 | 109658 | 4391 |
| 20 | 2 | 451430 | 112693 | 4520 |
| 20 | 5 | 443921 | 111069 | 4445 |
| 20 | 10 | 439961 | 110056 | 4407 |
| 20 | 20 | 437425 | 109399 | 4382 |
| Average | 451580 | 112945 | 4525 |  |

Table 2: Estimation of the expected makespan for the exponential testbed: Number of simulation runs required
less than 10,000 runs, but a $0.5 \%$ error requires more than 400,000 runs even for instances with small variability, and around $1,000,000$ for $c=0.5$.

In general, the need of more runs decreases with the number of machines, but remains relatively stable with respect to the number of jobs. This may speak for a compensation of the processing times of a job across the machines.

Additional experiments were carried out to establish the percentage error $(p)$ induced when the number of simulation runs is considered fixed. To do so, we estimate $\bar{C}_{\text {max }}$ for a random sequence using 100,000 simulation runs, and calculate a confidence interval for $\alpha=0.001$. By doing so, we are 'almost' sure than the makespan is contained in this interval, and then use the extreme values of this interval to obtain $p$. The results are shown in Table 3 and make clear that, while 100,000 simulation runs are an acceptable number for low variability, for moderate/high variability, this number of simulations leads to an average error over $10 \%$, and that, in some instances, this error is as high as $80 \%$.

The results show that it is difficult to be confident in the results obtained for the number of simulation runs employed in the literature. Estimation errors of less than \%1 require more than

|  |  | CV $=0.01$ | CV $=0.1$ | CV=0.2 | CV $=0.5$ | CV $=1.0$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 5 | 2 | 0.010 | 0.011 | 0.011 | 0.022 | 0.142 |
| 5 | 5 | 0.010 | 0.010 | 0.011 | 0.018 | 0.138 |
| 5 | 10 | 0.010 | 0.010 | 0.011 | 0.016 | 0.098 |
| 5 | 20 | 0.010 | 0.010 | 0.011 | 0.014 | 0.079 |
| 10 | 2 | 0.010 | 0.010 | 0.011 | 0.019 | 0.195 |
| 10 | 5 | 0.010 | 0.010 | 0.011 | 0.016 | 0.102 |
| 10 | 10 | 0.010 | 0.010 | 0.011 | 0.014 | 0.072 |
| 10 | 20 | 0.010 | 0.010 | 0.011 | 0.013 | 0.074 |
| 15 | 2 | 0.010 | 0.010 | 0.011 | 0.017 | 0.105 |
| 15 | 5 | 0.010 | 0.010 | 0.011 | 0.015 | 0.149 |
| 15 | 10 | 0.010 | 0.010 | 0.011 | 0.014 | 0.089 |
| 15 | 20 | 0.010 | 0.010 | 0.010 | 0.013 | 0.062 |
| 20 | 2 | 0.010 | 0.010 | 0.011 | 0.015 | 0.106 |
| 20 | 5 | 0.010 | 0.010 | 0.011 | 0.014 | 0.091 |
| 20 | 10 | 0.010 | 0.010 | 0.011 | 0.013 | 0.065 |
| 20 | 20 | 0.010 | 0.010 | 0.010 | 0.012 | 0.063 |
| Avg. | 0.010 | 0.010 | 0.011 | 0.015 | 0.102 |  |
| Max. | 0.010 | 0.011 | 0.011 | 0.026 | 0.815 |  |
| Min. | 0.010 | 0.010 | 0.010 | 0.012 | 0.037 |  |

Table 3: Percentual error in makespan estimation for $N=100,000$

200,000 runs for scenarios with medium/high variability. In addition, our method allows to know the accepted error of the estimations (in percentual terms) and therefore to assert the significance of the differences in the performance of solution procedures.

## 5 Comparison of heuristics

In this section, we carry out a computational study to establish the performance of different heuristics for the problem according to the procedure for the estimation of the expected makespan presented in the previous section. The heuristics tested are the following:

- The stochastic version of the NEH heuristic as described in Baker and Altheimer (2012). This heuristic is labelled SNEH.
- The stochastic version of the CDS/Talwar heuristic as described in Baker and Altheimer (2012). This heuristic is labelled SCDS/Talwar.
- The deterministic NEH heuristic applied using as data the mean processing times of the instances.
- The deterministic CDS/Talwar heuristic applied using as data the mean processing times of the instances.
- The deterministic NEH heuristic applied using as data the mean processing times of the instances, but using as initial order that given by the deterministic CDS/Talwar heuristic. This heuristic is labelles NEH-Talwar.

Note that, although the procedure for SNEH and SCDS/Talwar are identical to that of Baker and Altheimer (2012), the estimation of the expected makespan of the subsequences and that of the final sequence is carried out according to the procedure presented in Section 4 for $p=0.01$. Analogously, in order to estimate the expected makespan given by the deterministic heuristics (NEH, CDS/Talwar, and NEH-Talwar), the sequence obtained is evaluated following the aforementioned procedure.

The testbed presented in Section 3 is solved using the five heuristics presented above. Tables 4 to 7 show the results obtained for different values of $c$. Apart from the average values obtained by the estimated makespan for each one of the heuristic (labelled as Avg. in the tables), the average percentage increase of the makespan of each heuristic with respect to that of SNEH is presented (labelled as $\Delta$ in the tables).

In these tables, we do not give information on the time required for each heuristic. Note that the deterministic heuristic are nearly instantaneous for the problem sizes tested (e.g. NEH is less than 0.01 seconds for the biggest instances), although for our purposes, we have to evaluate the so-obtained sequence (something not required when applying it for a real problem). Regarding the times required for the stochastic heuristic, they depend obviously on the problem size and on the value of $c$, ranging from 1300 seconds for $c=0.01, n=20, m=20$ to 2000 seconds for $c=0.5, n=20, m=20$. In total, the computational workload of the experiments contained in the tables can be measured in weeks of CPU time.

Several comments can be done on the results obtained:

- With respect to the heuristics specifically designed for the stochastic problem, there are sig-
nificant differences in performance. This result contradicts those obtained by Baker and Altheimer (2012), who did not detect significant differences among them. It has to be noted that their way to estimate $E\left[C_{\max }\right]$ the solution is based on a fixed number of simulation runs and that the number that they employed $(100,000)$ has been proved to be unsufficient to establish consistent results for medium/large coefficient of variations. In addition, our testbed is of bigger size, a fact that may also explain some differences.
- The differences in performance between SNEH and SCDS/Talwar decrease with the variability of the testbed (from roughly $15 \%$ for $c=0.01$ to about $6 \%$ for $c=0.5$ ), but still are substantial for relatively large coefficient of variations. The explanation may lie in the fact that Talwar's rule is optimal for the exponential distribution, whose coefficient of variation is 1 , and therefore its performance improves for instances where $c$ is closer to that value.
- With respect to the deterministic heuristic tested, the best performance correspond to the NEH. It is interesting to note that CDS/Talwar is suppose to incorporate some stochastic considerations in their ordering, however these do not pay off, either as a simple heuristic, or as a starting order for the NEH heuristic.
- When comparing SNEH and NEH it may be seen that the differences are very small. In some cases, these differences are below the error accepted for $p(1 \%)$. The conclusion is that, for a realistic ranges of variability in the flowshop, the application of the NEH to the mean processing times gives an extremely good estimation of the performance of their stochastic counterpart. If we note that SNEH is highly CPU-intensive, requiring a high number of simulations of all subsequences for all steps, it has to be questioned whether the effort in running SNEH pays off.
- Although there is no clear pattern, it seems that the differences in performance of all the heuristics decrease with the number of machines, a fact for which we do not have at the
moment a fully convincing explanation.

Table 4: Comparative results of the different heuristics for $c=0.01$.

|  |  | SNEH | SCDS/Talwar |  | NEH | CDS/Talwar |  |  |  | NEH-Talwar |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: |
| $n$ | $m$ | Avg. | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. | $\Delta$ |  |  |
| 5 | 2 | 309.848 | 348.360 | 13.056 | 309.978 | 0.041 | 348.367 | 13.059 | 310.171 | 0.103 |  |  |
| 5 | 5 | 477.448 | 550.888 | 15.905 | 477.552 | 0.024 | 550.878 | 15.903 | 477.373 | -0.086 |  |  |
| 5 | 10 | 795.812 | 891.632 | 12.574 | 795.833 | 0.004 | 891.644 | 12.576 | 798.258 | 0.301 |  |  |
| 5 | 20 | 1372.339 | 1482.957 | 8.063 | 1374.269 | 0.148 | 1482.933 | 8.061 | 1374.786 | 0.167 |  |  |
| 10 | 2 | 571.929 | 632.638 | 11.423 | 572.181 | 0.046 | 632.649 | 11.426 | 572.615 | 0.136 |  |  |
| 10 | 5 | 720.947 | 866.235 | 20.268 | 721.875 | 0.119 | 866.224 | 20.266 | 716.499 | -0.575 |  |  |
| 10 | 10 | 1049.066 | 1246.054 | 18.884 | 1053.062 | 0.384 | 1246.057 | 18.885 | 1052.709 | 0.348 |  |  |
| 10 | 20 | 1673.640 | 1902.705 | 13.662 | 1676.495 | 0.169 | 1902.713 | 13.662 | 1681.827 | 0.490 |  |  |
| 15 | 2 | 801.950 | 857.986 | 6.943 | 802.028 | 0.010 | 857.982 | 6.943 | 805.056 | 0.406 |  |  |
| 15 | 5 | 1008.644 | 1206.445 | 19.737 | 1008.840 | 0.017 | 1206.441 | 19.736 | 1013.747 | 0.536 |  |  |
| 15 | 10 | 1309.681 | 1574.132 | 20.336 | 1312.544 | 0.223 | 1574.124 | 20.335 | 1324.418 | 1.134 |  |  |
| 15 | 20 | 1965.180 | 2261.390 | 15.063 | 1981.237 | 0.833 | 2261.389 | 15.063 | 1972.786 | 0.375 |  |  |
| 20 | 2 | 1085.429 | 1155.937 | 6.582 | 1085.508 | 0.008 | 1155.929 | 6.582 | 1088.635 | 0.309 |  |  |
| 20 | 5 | 1224.466 | 1486.417 | 21.419 | 1228.891 | 0.368 | 1486.406 | 21.418 | 1240.408 | 1.305 |  |  |
| 20 | 10 | 1601.085 | 1939.413 | 21.186 | 1609.659 | 0.548 | 1939.435 | 21.187 | 1615.548 | 0.926 |  |  |
| 20 | 20 | 2258.262 | 2658.262 | 17.759 | 2269.505 | 0.501 | 2658.259 | 17.759 | 2274.953 | 0.746 |  |  |
|  |  |  |  | 15.179 |  | 0.215 |  | 15.179 |  | 0.414 |  |  |

## 6 Conclusions

In this paper, we have addressed the problem of scheduling jobs in a flowshop when their processing times adopt a given distribution. Existing literature for the problem reveals that optimal solutions can be found only for very specific cases, so some heuristics with similar performance have been proposed for the general case. We first focus on the critical issue of estimating the expected makespan of a sequence, and found that, for instances with a medium/large variability (expressed as the coefficient of variation of the processing times of the jobs), the number of samples (simulation runs) used in the literature to estimate the expected makespan of a sequence may not be sufficient to derive conclusive results. We propose a procedure with a variable number of iterations that ensures that the error in the estimation of the expected makespan is bounded (with a very high probability) within a small percentage. Using this procedure, we test the main

Table 5: Comparative results of the different heuristics for $c=0.1$.

|  |  | SNEH | SCDS/Talwar | NEH |  |  |  | CDS/Talwar |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $n$ | $m$ | Avg. | Avg. | $\Delta$ | Avg. | $\Delta$ | NEH-Talwar |  |  |  |
| 5 | 2 | 291.782 | 325.107 | 11.687 | 293.834 | 0.772 | 325.137 | 11.698 | 295.252 | 1.175 |
| 5 | 5 | 475.155 | 532.005 | 12.307 | 478.174 | 0.631 | 532.057 | 12.317 | 476.358 | 0.225 |
| 5 | 10 | 782.803 | 860.982 | 10.386 | 788.699 | 0.763 | 862.039 | 10.538 | 790.445 | 0.971 |
| 5 | 20 | 1348.199 | 1439.980 | 6.817 | 1358.467 | 0.787 | 1440.195 | 6.832 | 1355.911 | 0.557 |
| 10 | 2 | 529.584 | 587.962 | 11.609 | 539.093 | 1.782 | 587.953 | 11.609 | 539.663 | 1.955 |
| 10 | 5 | 722.250 | 843.684 | 16.890 | 733.248 | 1.535 | 843.613 | 16.879 | 733.328 | 1.524 |
| 10 | 10 | 1071.549 | 1233.335 | 15.202 | 1086.796 | 1.465 | 1233.383 | 15.206 | 1087.909 | 1.536 |
| 10 | 20 | 1697.238 | 1873.897 | 10.362 | 1715.366 | 1.079 | 1873.586 | 10.344 | 1722.355 | 1.489 |
| 15 | 2 | 746.468 | 806.670 | 7.981 | 754.396 | 1.071 | 806.746 | 7.992 | 764.054 | 2.332 |
| 15 | 5 | 993.123 | 1171.510 | 18.038 | 1010.630 | 1.783 | 1171.513 | 18.041 | 1019.869 | 2.717 |
| 15 | 10 | 1352.968 | 1555.070 | 15.059 | 1371.496 | 1.386 | 1555.117 | 15.062 | 1378.176 | 1.874 |
| 15 | 20 | 2033.541 | 2269.854 | 11.609 | 2069.860 | 1.795 | 2271.566 | 11.697 | 2061.274 | 1.359 |
| 20 | 2 | 1005.958 | 1079.875 | 7.376 | 1016.039 | 0.998 | 1079.903 | 7.378 | 1024.590 | 1.865 |
| 20 | 5 | 1221.532 | 1440.011 | 17.865 | 1245.929 | 1.994 | 1440.785 | 17.929 | 1261.137 | 3.217 |
| 20 | 10 | 1647.583 | 1919.185 | 16.478 | 1682.258 | 2.108 | 1919.258 | 16.482 | 1687.571 | 2.416 |
| 20 | 20 | 2351.119 | 2686.962 | 14.339 | 2394.679 | 1.851 | 2686.998 | 14.341 | 2397.252 | 1.959 |
|  |  |  |  | 12.750 |  | 1.362 |  | 12.771 | 1.698 |  |

Table 6: Comparative results of the different heuristics for $c=0.2$.

|  |  | SNEH | SCDS/Talwar |  | NEH |  | CDS/Talwar |  | NEH-Talwar |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $n$ | $m$ | Avg. | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. |  |
| 5 | 2 | 322.603 | 352.795 | 9.441 | 325.613 | 1.004 | 352.802 | 9.441 | 327.488 | 1.52 |
| 5 | 5 | 562.439 | 613.089 | 9.259 | 566.252 | 0.688 | 613.133 | 9.264 | 564.659 | 0.38 |
| 5 | 10 | 934.795 | 1007.518 | 8.112 | 941.541 | 0.743 | 1009.800 | 8.384 | 941.668 | 0.74 |
| 5 | 20 | 1601.737 | 1697.290 | 5.986 | 1615.790 | 0.895 | 1697.146 | 5.978 | 1612.897 | 0.68 |
| 10 | 2 | 581.957 | 641.934 | 10.700 | 599.805 | 3.022 | 641.875 | 10.688 | 599.746 | 3.10 |
| 10 | 5 | 871.709 | 988.245 | 13.382 | 886.699 | 1.719 | 988.071 | 13.360 | 887.356 | 1.75 |
| 10 | 10 | 1334.107 | 1495.436 | 12.168 | 1352.889 | 1.447 | 1495.574 | 12.177 | 1354.049 | 1.50 |
| 10 | 20 | 2116.182 | 2291.745 | 8.248 | 2138.487 | 1.065 | 2291.817 | 8.252 | 2146.237 | 1.43 |
| 15 | 2 | 817.669 | 888.258 | 8.541 | 836.351 | 2.297 | 888.129 | 8.528 | 849.235 | 3.81 |
| 15 | 5 | 1197.086 | 1373.755 | 14.815 | 1222.018 | 2.101 | 1373.915 | 14.828 | 1231.843 | 2.93 |
| 15 | 10 | 1693.012 | 1899.637 | 12.271 | 1723.999 | 1.839 | 1899.451 | 12.259 | 1730.821 | 2.24 |
| 15 | 20 | 2589.667 | 2834.702 | 9.461 | 2636.059 | 1.798 | 2836.892 | 9.551 | 2627.271 | 1.45 |
| 20 | 2 | 1093.098 | 1185.651 | 8.490 | 1119.084 | 2.378 | 1185.546 | 8.480 | 1134.384 | 3.79 |
| 20 | 5 | 1467.477 | 1688.517 | 15.032 | 1502.591 | 2.381 | 1689.583 | 15.106 | 1521.990 | 3.69 |
| 20 | 10 | 2072.442 | 2350.040 | 13.380 | 2123.447 | 2.468 | 2350.523 | 13.403 | 2128.186 | 2.68 |
| 20 | 20 | 3025.734 | 3381.442 | 11.808 | 3082.084 | 1.859 | 3381.093 | 11.796 | 3084.835 | 1.94 |
| Avg |  |  |  | 10.693 |  | 1.731 |  | 10.718 |  | 2.10 |

Table 7: Comparative results of the different heuristics for $c=0.5$.

|  |  | SNEH | SCDS/Talwar |  | NEH |  | CDS/Talwar |  | NEH-Talwar |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $n$ | $m$ | Avg. | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. |  |
| 5 | 2 | 744.176 | 776.915 | 4.380 | 748.874 | 0.657 | 776.746 | 4.365 | 752.611 | 1.1 |
| 5 | 5 | 1529.894 | 1595.708 | 4.422 | 1537.664 | 0.520 | 1595.128 | 4.392 | 1533.510 | 0.2 |
| 5 | 10 | 2726.438 | 2842.900 | 4.478 | 2741.956 | 0.596 | 2848.047 | 4.689 | 2739.926 | 0.5 |
| 5 | 20 | 4807.027 | 4982.471 | 3.657 | 4836.312 | 0.613 | 4986.449 | 3.744 | 4834.410 | 0.5 |
| 10 | 2 | 1411.348 | 1496.158 | 6.206 | 1446.847 | 2.448 | 1496.660 | 6.243 | 1446.062 | 2.4 |
| 10 | 5 | 2579.011 | 2764.617 | 7.170 | 2610.984 | 1.251 | 2763.582 | 7.132 | 2613.878 | . 3 |
| 10 | 10 | 4384.309 | 4678.766 | 6.734 | 4433.741 | 1.153 | 4684.903 | 6.870 | 4433.034 | 1.1 |
| 10 | 20 | 7312.968 | 7656.029 | 4.647 | 7357.016 | 0.616 | 7659.549 | 4.700 | 7373.893 | 0.8 |
| 15 | 2 | 1979.146 | 2095.786 | 5.837 | 2021.776 | 2.165 | 2095.632 | 5.814 | 2042.227 | 3.1 |
| 15 | 5 | 3683.235 | 3989.336 | 8.347 | 3737.412 | 1.467 | 3990.114 | 8.363 | 3756.730 | 2.0 |
| 15 | 10 | 5793.048 | 6186.216 | 6.805 | 5861.646 | 1.184 | 6189.823 | 6.863 | 5878.701 | 1.4 |
| 15 | 20 | 9561.606 | 10073.656 | 5.356 | 9661.007 | 1.039 | 10078.809 | 5.412 | 9646.360 | 0.8 |
| 20 | 2 | 2647.731 | 2821.413 | 6.605 | 2711.862 | 2.447 | 2819.596 | 6.535 | 2743.518 | 3.6 |
| 20 | 5 | 4527.154 | 4910.429 | 8.443 | 4590.192 | 1.394 | 4914.857 | 8.541 | 4632.402 | 2.3 |
| 20 | 10 | 7330.081 | 7883.050 | 7.532 | 7445.274 | 1.580 | 7882.278 | 7.528 | 7446.577 | 1.5 |
| 20 | 20 | 11624.512 | 12411.446 | 6.818 | 11730.277 | 0.911 | 12403.771 | 6.757 | 11745.989 | 1.0 |
| Avg. |  |  |  | 6.090 |  | 1.252 |  | 6.122 |  | 1.5 |

Table 8: Comparative results of the different heuristics for $c=1.0$.

|  |  | SNEH | SCDS/Talwar |  | NEH |  | CDS/Talwar |  | NEH-Talwar |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $n$ | $m$ | Avg. | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. | $\Delta$ | Avg. |  |
| 5 | 2 | 6266.040 | 6341.156 | 1.201 | 6290.683 | 0.417 | 6350.017 | 1.321 | 6303.186 |  |
| 5 | 5 | 14794.404 | 14998.899 | 1.432 | 14825.357 | 0.213 | 15011.184 | 1.529 | 14822.093 |  |
| 5 | 10 | 29495.484 | 29905.531 | 1.464 | 29539.868 | 0.158 | 29919.130 | 1.506 | 29527.008 |  |
| 5 | 20 | 55811.357 | 56727.477 | 1.631 | 56006.588 | 0.345 | 56705.085 | 1.604 | 56013.200 |  |
| 10 | 2 | 13280.157 | 13505.036 | 1.777 | 13399.913 | 0.841 | 13497.332 | 1.710 | 13391.054 |  |
| 10 | 5 | 27912.319 | 28581.464 | 2.392 | 28074.882 | 0.575 | 28571.105 | 2.334 | 28096.549 |  |
| 10 | 10 | 53943.962 | 55067.908 | 2.081 | 54019.570 | 0.141 | 55114.819 | 2.151 | 54034.985 |  |
| 10 | 20 | 98731.447 | 100432.899 | 1.710 | 99116.154 | 0.407 | 100587.218 | 1.857 | 99022.914 |  |
| 15 | 2 | 18628.113 | 18968.050 | 1.798 | 18732.248 | 0.565 | 18958.325 | 1.754 | 18819.655 |  |
| 15 | 5 | 43253.615 | 44324.220 | 2.494 | 43325.129 | 0.174 | 44268.685 | 2.381 | 43416.154 |  |
| 15 | 10 | 75215.521 | 77090.704 | 2.488 | 75648.411 | 0.556 | 77170.600 | 2.581 | 75667.356 |  |
| 15 | 20 | 139988.391 | 142683.498 | 1.916 | 140496.048 | 0.353 | 142720.286 | 1.949 | 140481.886 |  |
| 20 | 2 | 25754.094 | 26255.813 | 1.981 | 25897.427 | 0.584 | 26234.195 | 1.891 | 26025.904 |  |
| 20 | 5 | 53797.160 | 55212.357 | 2.618 | 53938.933 | 0.273 | 55249.351 | 2.688 | 54102.808 |  |
| 20 | 10 | 101187.069 | 103972.650 | 2.752 | 101715.555 | 0.530 | 103915.201 | 2.684 | 101785.109 |  |
| 20 | 20 | 179560.039 | 183853.265 | 2.433 | 180249.784 | 0.402 | 184245.565 | 2.638 | 180325.990 |  |
| Avg. |  |  |  | 2.010 |  | 0.408 |  | 2.036 |  |  |

heuristic proposed in the literature in a larger testbed and found significant differences in their performance, in contrast with existing studies. We also found that the deterministic counterpart of the most efficient heuristic for the stochastic problem performs extremely well for most settings, which indicates that (at least within the limitations of our study), a practical way to solve the $F m|p r m u| E\left[C_{\max }\right]$ is to simplify it to its deterministic version.

Several issues lie ahead as future research lines. First of all, the computational analysis can be enhanced by introducing different families of distribution, higher coefficients of variation for the lognormal distribution, and bigger problem instances. However, it is to note that some families of distributions widely applied in theory due to their good properties (such as e.g. the exponential) do not match well with the processing times distributions found in practice. Analogously, very large coefficient of variations are not the norm in industry, as discussed before. Finally, there are severe computational limitations for using bigger testbeds, as the experiments carried out in this paper already amount for weeks of CPU time.

A more fruitful research line may be to check whether the results obtained by the best stochastic heuristic (SNEH) is close to the optimal values, or not. Although this neccessarily has to be done for small instances, it would be interesting to confirm the results by Baker and Altheimer (2012), who indicate that SNEH obtains close-to-optimal solutions. If this was not the case, ground for new approximate solutions for the problem may be re-opened.

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